# REGIONAL DISCRIMINATION OF CHANGE IN RUNOFF

by

Viboon Nimmannit and Hubert J. Morel-Seytoux

November 1969



HYDROLOGY PAPERS
COLORADO STATE UNIVERSITY
Fort Collins, Colorado

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# RELATION OF HYDROLOGY PAPER NO. 37 TO RESEARCH PROGRAM: "HYDROLOGY OF WEATHER MODIFICATION"

The present study is part of a more comprehensive project which has as one of its objectives the development of methods of evaluation of atmospheric water resources programs. Correlatively the application of the methods to a variety of basins forms a basis for selection of suitable watersheds, basins or regions.

Several approaches were pursued. This report discusses one of them.

Several other approaches were previously described in Hydrology Papers 22, 34, and 36 (see back inside cover for complete reference).

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#### ABSTRACT

The object of this study is to find answers to the following questions:

What is the appropriate statistical test for a regional target-control technique of evaluation?

What is a suitable method for reduction of an originally large number of variables?

Which of the Upper Basin of the Colorado River or the San Juan Mountains is a more suitable area of operations, if the effectiveness of precipitation management is to be detected as quickly as possible?

The results of this research study show:

- 1. The  $T^2$ -test is the appropriate test for multiple target-control technique of evaluation.
- The canonical analysis is the suitable method for the reduction of a large number of original variables.
- 3. The Upper Basin of the Colorado River is preferable under the assumption of an equal percentage of increase in runoff. However, if the percentage increase in the southern area is at least 1.2 times as large as in the northern area (and recent publications suggest that this ratio is probably around 3) then the southern area is preferable.

Based on the  $T^2$ -test, the minimum number of years for detecting an increase of 10 percent in spring runoff means are three years in the Upper Basin of the Colorado River, and four years in the San Juan Mountains.

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#### Chapter I

#### INTRODUCTION

1.1 Motivation of study. As interference with nature is accelerating [1,2,3] there is a need for early detection of direct or side effects of man's actions. Because of the rapid pace of development [3, 4,5] it is important to develop techniques that will display the effect of any given practice on water resources availability and distribution at the earliest possible time. For large scale field research, the availability of an efficient and regionally representative test would reduce the duration of experiments required to attain conclusive results and therefore costs, and provide a basis for managerial decision at an earlier stage, without additional observations. The decision may be to stop a project earlier when it becomes apparent, based on real time analysis of data, that the objectives cannot be achieved in the planned time. Better, pre-experiment data simulation would permit to assess the chances of being in that unfortunate situation as a function of a range of values of the suspected or hoped for change. Useful charts can be drawn in terms of the parameters, (magnitude of change, basin characteristics, etc.) for first stage planning.

The techniques which are described in this paper could be used for detection of the effects of watershed management of any origin upon water supply. They could be used to determine the effect of urbanization on the local hydrology, to detect when such urbanization has created a significant change that calls for reappraisal of the protective designs, e.g., flood control, etc. In other words, they are quite general. To a certain degree the techniques will indeed be discussed in a general abstract form, but their practical applicability will be demonstrated with a very special and very important application in mind.

The Bureau of Reclamation will most probably initiate in the fall of 1970 a pilot project of massive cloud seeding operations, covering some 4000 square miles within the state of Colorado. It will be the primary purpose of this paper to establish as accurately as possible how long it will take to detect a regional hydrologic change and to attribute it with little risk of error to the cloud seeding operations. To understand this practical illustration of the technique some knowledge of the geographic and hydrologic features of the region, of the water situation and of the plans of the Bureau of Reclamation is a prerequisite. The purpose of the following sections is to provide this background information.

- 1.2 Geographic and hydrologic setting. The Colorado River begins high in the snow-capped Rocky Mountains of north central Colorado, flows nearly 1,400 miles southwest, and empties into the Gulf of California in Mexico far to the south. It drains a vast area of 244,000 square miles, 242,000 square miles in the United States -- one-twelfth of the area of Continental United States -- and 2,000 square miles in northern Mexico. The basin from Wyoming to below the Mexican border is some 900 miles long and varies in width from about 300 miles in the upper section to 500 miles in the lower section. It is bounded on the north and east by the Continental Divide in the Rocky Mountains, on the west by the Wasatch Range, and on the southwest by the San Jacinto Mountains, a range of the Sierra Nevada Mountains. The area, larger than the states of New York, Pennsylvania, and New Jersey combined, above Lee Ferry, Arizona, is known as the Upper Colorado River Basin (Fig. 1). This area is the source of the greatest part of water reaching the Colorado River. The upper portion of this basin in Wyoming and Colorado is a mountainous plateau, 5,000 to 8,000 feet in altitude, marked by broad rolling valleys, deep canyons, and intersecting mountain ranges. Climatologically, the Colorado River Basin has heavy precipitation on the high peaks of the Rockies and truly desert conditions with little rain in the southern area around Yuma, Arizona. Extremes of temperatures in the basin range from 50° below zero to 130° above zero degree Fahrenheit. Development and utilization of resources in this arid land depend on the availability of water. Crops must be irrigated; cattle on the vast ranges must be partially fed from hay produced on irrigated land; towns and cities must be located within distance of dependable domestic and municipal water supplies, and mining and many other industries depend, to an extent, on the availability of hydroelectric power [1].
- 1.3 The water resources outlook. The U.S. Geological Survey estimates total water demand in the United States was 280 billion gallons per day (314 million acre-feet per year) in 1960. As a point of comparison let us note that the average annual flow of the biggest river in the United States, the Mississippi, is 440 maf and that of the Upper Colorado is about 14 maf. The U.S.G.S. estimates the total water demand for the U.S. will be 600 billion gallon per day (672 million acre-feet per year) by 1980. In 1960 the demand in the Western States alone was estimated at 125 billion gallons per day (140 million

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Fig. 1 The Upper Colorado River Basin (after Upper Colorado River Commission)

acre-feet per year) and for 1980 at 190 billion gallons per day (213 million acre-feet per year). The lower percentage of demand growth for the Western States reflects different demands of industry in the East and agriculture in the West. Because rainfall is low in the Western States, the conservation use must be greater than in the East and Midwest. Municipal or domestic use has first priority in the West, with irrigation second. It is estimated the 44,000,000 population of the Western States in 1960 will expand to more than 100,000,000 by the year 2000 [2].

From the population figures given above, it is obvious much more water will be needed in the near future. So, the question one must answer is, "What can be used as sources for additional water to alleviate the shortages?" Several agencies, such as, the Bureau of Reclamation [5], the Upper Colorado River Commission [4], and the Committee on Water of the National Research Council [5], feel cloud seeding, to augment the precipitation amount in the Upper Colorado River Basin, may become a partial solution to the recurrent water shortage.

1.4 Precipitation management operations and plans. An important experimental cloud seeding operation is being conducted near Climax, Colorado, by Colorado State University under sponsorship of the National Science Foundation. These experiments are designed to show quantitative change in precipitation by cloud seeding and to determine criteria for optimum seeding conditions.

The most favorable conditions for cloud seeding are in regions where moist winds blow more or less constantly up the slopes of the mountains. Cloud seeding involves artificial introduction of tiny particles into clouds so that moisture can depose around each of the nuclei to form a crystal heavy enough to fall to the ground. Among nuclei that have been used experimentally in cloud seeding operations are solid carbon dioxide, silver iodide, water spray, and carbon black. To date, the greatest number of cloud-seeding attempts have been made by using silver iodide generators operated on the ground. However, seeding operations using aircraft flown directly over cloud layers have demonstrated that this technique may be more effective [6].

In 1968, the Bureau of Reclamation adopted a plan to start pilot programs for weather modification operations in the Upper Colorado River Basin (Fig. 1), and two regions were selected for this purpose [7]. The first was the Upper Basin of the Colorado River\*, which will for brevity be referred to in this study as the Northern Project area (Fig. 2). The second area was the San Juan Mountains region referred to as the Southern Project area (Fig. 3). Since the initiation of this study, the plans of the Bureau were modified. Currently [8] only one area is considered: the Southern area. Nevertheless, because they had already been calculated, the results for the Northern area are also reported.

1.5 Objective of study and approach. The primary objective was to develop an appropriate and efficient methodology that can be used to demonstrate the effectiveness of cloud seeding in each project region. In order to achieve this, a multivariate analysis of geographically well distributed stations in each region is carried out. These stations are referred to as targets. Variables used in this study are spring runoffs. The spring runoff of a station is defined here as the average flow, in cubic feet per second, of that station during the spring months. Because this flow is substantially contributed by winter snow, it can be regarded as an indirect measure of the effect of weather modification. However, because of the lack of a precise date for the start of snow melting, two different time intervals will be used for spring months. The first interval will be composed of four months: April, May, June and July; the second of six months: March, April, May, June, July and August.

Because the use of controls, which are the stations free from the effect of weather modification, is a well proven means of making tests more effective, (9), it also will be utilized in this study. An area between the Northern and Southern Project areas has been selected (Fig. 4) to serve as the control area.

<sup>\*</sup>The reader is warned for possible confusion. In this paper the expression "Upper Colorado River Basin" refers to the Colorado Basin above Lee's Ferry. On the other hand, the expression "Upper Basin of the Colorado River" refers to a much smaller drainage basin including the main stem of the Colorado close to its source and a few tributaries. The limits of that basin are shown on Fig. 2.

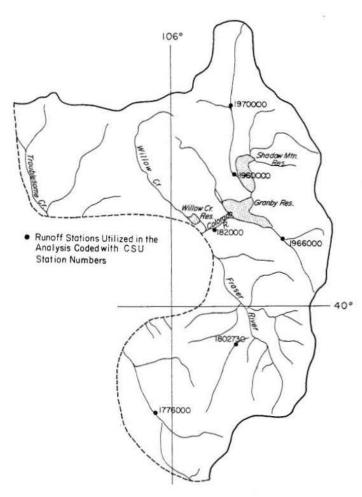


Fig. 2 General configuration of and location of gages within the Upper Basin of the Colorado River

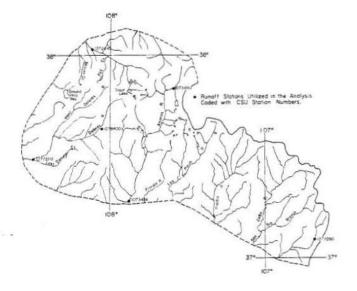


Fig. 3 General configuration of and location of gages within the Colorado River Basin Pilot Project area

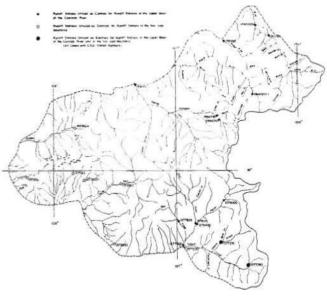


Fig. 4 General configuration of and location of gages within the Colorado River Basin Pilot Project control area

For brevity, the following symbols will be employed:

- N-4: 4-month runoff series in the northern
- target region,
  N-6: 6-month runoff series in the northern target region,
- CN-4: 4-month runoff series in the northern
- control region,
  CN-6: 6-month runoff series in the northern control region,
- S-4: 4-month runoff series in the southern
- target region, S-6: 6-month runoff series in the southern
- CS-4: target region, 4-month runoff series in the southern
- CS-6: control region, 6-month runoff series in the southern control region.
- N-CN-4: the combination of N-4 and CN-4,
- N-CN-6: the combination of N-6 and CN-6,
- S-CS-4: the combination of S-4 and CS-4,
- S-CS-6: the combination of S-6 and CS-6.

In applying theories of statistics to an engineering problem, it is necessary to assume certain properties of the variables. The assumptions made in this study are:

- a) The observations of runoff follow a multivariate normal distribution.
- b) The estimated means in both target and control areas from the period before seeding are essentially equal to the population values.
- c) After seeding the means in the target areas will change but the means in the control areas will remain unchanged.
- d) The covariance matrix of the target and control variables is the same for both periods before and after seeding.

The above assumptions are required in this study because of the difficulty in developing the theoretical distribution of the test criterion otherwise. In dealing with more than two variables, the knowledge of distributions, except that of the normal distribution, are not sufficiently developed [10]. So, even though it is rather obvious the assumptions made here will be violated to some degree in reality, they are practically as good as one can make with the present state of statistical knowledge.

From the work of Ref. [9], it is found that the  $\chi^2$ -test which is based on the population values, and the conditional Student's t-test which is based on the sample values, give very closely the same results for sample sizes around 30. Thus, for convenience in handling the mathematics, the population values are assumed to be known here and this assumption appears justified. Also, all the observations of runoff station used in this study have been plotted on normal probability paper. If the runoff were exactly distributed as a normal variate, all the observations would fall exactly on a straight line. The actual observations did not in any case deviate appreciably from a straight line. The assumption of normality may therefore be entertained for these data.

Based on the above assumptions, a  $T^2$ -statistic is obtained [11,12]. The minimum number of years,  $N^*$ , to detect the expected increase can be obtained [11] from the formula,

$$N^* = \frac{\tau^2}{\mu^* V^{-1} \mu}$$
 , (1)

where  $\tau^2$  is the noncentrality parameter (it is a measure of the amount of deviation from being central which is the case when the variables under study have means zero),

 $\underline{\mu} = \underline{\mu}^* - \underline{\mu}_0$ ,  $\underline{\mu}^*$  is the runoff mean vector for the seeded period, and  $\underline{\mu}_0$  is the runoff mean vector for the non-seeded period,

 $\underline{\mu}$ ' is the transpose of  $\underline{\mu}$  , and

 $\underline{V}^{-1}$  is the inverse of the covariance matrix of runoff variables, V.

In Chapter II, most approaches used to detect the effectiveness of weather modification by other investigators are summarized. The theoretical concepts of the principal component analysis, the canonical analysis, and the  $T^2$ -statistic are the main subjects of Chapter III. Chapters IV and V deal with data assembly, analysis of data, and results.

The study led to two major conclusions, one of general theoretical interest and the second of practical significance for the plans of the Bureau:

- a) Canonical analysis coupled with the multivariate  ${\bf T}^2$ -test provides an effective technique of detection of a suspected regional hydrologic change and.
- b) Assuming a 10% uniform increase in runoff by precipitation management 3 and 4 years only are required for significant evaluation for the Upper Basin of the Colorado and the San Juan Mountains, respectively.

#### Chapter II

#### REVIEW OF PREVIOUSLY USED TESTS

The statistical content of this chapter is not new. The material here is provided for the sake of convenience to a reader whose statistical background is that of the average engineer. A statistician can bypass this chapter without detrimental effect to the continuity and understanding of this paper.

In this chapter the statistical tests, which have been employed by other investigators for detecting the effectiveness of weather modification, will be presented. The literature is further discussed in Ref. 12. Because all tests are concerned with the expected increase in the means of either runoff or precipitation during the seeded period, the hypotheses for all tests can be stated as:

 $\rm H_{o}$  (null hypothesis) - there is no increase in the mean of the hydrologic variable during the seeded period,

 ${\rm H}_{\stackrel{\mbox{\scriptsize a}}{a}}$  (alternate hypothesis) - there is an increase in the mean.

2.1 Target sample u-test. Let  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n_1}$ , be  $n_1$  observations of a hydrologic variable for the nonseeded period, and  $q_{21}$ ,  $q_{22}$ , ...,  $q_{2n_2}$  be  $n_2$  observations for the seeded period of a target watershed. When  $n_1$  is large the mean and variance of the series  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n_1}$  can be considered to be the population mean and population variance. Assuming the variance of the seeded period is the same as the non-seeded period, the test statistic is [13]

$$u_o = \frac{\overline{q}_2 - \mu_1}{\sigma_1 / \sqrt{n_2}} \quad ,$$

where  $\mathbf{u}_{_{\mathbf{O}}}$  is normally distributed with mean o and variance 1

$$\overline{q}_2 = \frac{1}{n_2} \quad \sum_{i=1}^{n_2} \quad q_{2i}$$

$$\mu_1 = \frac{1}{n_1} \sum_{i=1}^{n_1} q_{1i}$$

$$\sigma_1^2 = \frac{1}{n_1} \sum_{i=1}^{n_1} (q_{1i} - \mu_1)^2 .$$

The null hypothesis,  $\rm H_{\odot}$ , will be accepted at a 5% level of significance if  $\rm u_{\odot}$  has a value less than 1.645. That is, there is no increase in the mean. On the contrary, if  $\rm u_{\odot}$  is greater than 1.645

the alternative hypothesis,  $\rm H_a$ , will be accepted at a 5% level of significance. The use of this test can be found in References [9] and [14]. South Fork San Joaquin, California, was the target basin for the study in Reference [9]. There were 15 years of seeded record, and 29 years of non-seeded record. The apparent percentage increase in the mean of the seasonal runoff for the seeded period was about 10%. By the use of the target sample u-test it was found that  $\rm u_0$  = 1.20. This shows that the target sample u-test was not powerful enough to detect the increase in mean value in the order of 10% of the old mean.

2.2 Target two-sample t-test. This test does not require knowledge of population parameters. Let  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n_1}$  and  $q_{21}$ ,  $q_{22}$ , ...,  $q_{2n_2}$  be  $n_1$  and  $n_2$  observations for the non-seeded and seeded periods of a target watershed.

Assuming the variances of the non-seeded and seeded periods are equal, the test statistic [15]

$$t_o = \frac{\overline{q}_2 - \overline{q}_1}{s\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

is distributed as t-distribution with  $n_1 + n_2 - 2$  degrees of freedom, where:

$$\overline{q}_1 = \frac{1}{n_1} \sum_{i=1}^{n_1} q_{1i}$$

$$\overline{q}_2 = \frac{1}{n_2} \sum_{i=1}^{n_2} q_{2i}$$

5

$$s^2 = \frac{\sum\limits_{i=1}^{n_1} (q_{1i} - \overline{q}_1)^2 + \sum\limits_{i=1}^{n_2} (q_{2i} - \overline{q}_2)^2}{(n_1 - 1) + (n_2 - 1)}$$

The use of this test can be found in References [8], [14], [16], [17], [18], [19], [20], [21], [22], [23], and [24]. The value of the t-statistic was also computed for South Fork San Joaquin [9] from the same set of data used in computing the target sample  $\mathbf{u}_{0}$ . The computed t-statistic has the value of 0.89. So, again no significant increase was concluded. The target two-sample t-test, and the target sample u-test therefore can be considered to be insufficiently powerful tests for studies of this nature.

2.3 Target-control  $\chi^2$ -test. The detectability of the test can be improved by the use of a control [9]. This can be done by comparing sets of hydrologic data of non-seeded and seeded periods for the target watershed with those for an unseeded control watershed located in the vicinity of the target area.

Let  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n_1}$  and  $q'_{11}$ ,  $q'_{12}$ , ...,  $q'_{1n_1}$  be  $n_1$  observations for the period prior to seeding of the target and control watersheds respectively. Also, let  $n_2$  observations for the seeded period in the target be denoted by  $q_{21}$ ,  $q_{22}$ , ...,  $q'_{2n_2}$ , and those in the control by  $q'_{21}$ ,  $q'_{22}$ , ...,  $q'_{2n_2}$ ,

When the length of record before seeding is long enough, the estimated statistics of the target and control can be assumed to be the population values. Assuming the variables in the target and control are bivariate normally distributed, then the test statistic [14]:

$$\chi_{0}^{2} = \frac{n_{2}}{1-\rho^{2}} \left\{ \left| \frac{\overline{q_{2}} - \mu_{1}}{\sigma} \right|^{2} - 2\rho \frac{(\overline{q_{2}} - \mu_{1})(\overline{q_{2}} - \mu_{1}')}{\sigma \sigma'} + \left| \frac{\overline{q_{2}} - \mu_{1}'}{\sigma'} \right|^{2} \right\}$$

is distributed as Chi-square distribution with two degrees of freedom, where

ρ is the population coefficient of correlation between the target and control for the nonseeded period, given by

$$\rho = \frac{\sum_{i=1}^{n_1} (q_{1i} - \mu_1) (q'_{1i} - \mu'_1)}{\left[\sum_{i=1}^{n_1} (q_{1i} - \mu_1)^2 \sum_{i=1}^{n_1} (q'_{1i} - \mu'_1)^2\right]^{\frac{1}{2}}}$$

$$u_1 = \frac{1}{n_1} \sum_{i=1}^{n_1} q_{1i}$$

$$\mu_{1}' = \frac{1}{n_{1}} \sum_{i=1}^{n_{1}} q_{1i}'$$

$$\overline{q}_2 = \frac{1}{n_2} \sum_{i=1}^{n_2} q_{2i}$$

$$\overline{q'_2} = \frac{1}{n_2} \sum_{i=1}^{n_2} q'_{2i}$$

$$\sigma = \sqrt{\frac{1}{n_1} \sum_{i=1}^{n_1} (q_{1i} - \mu_1)^2}$$

$$\sigma' = \sqrt{\frac{1}{n_1} \sum_{i=1}^{n_1} (q_{1i}' - \mu_1')^2} .$$

This test has been used in References [9] and [14].

With the use of Merced River at Pohono Bridge as a control runoff station for the target, South Fork San Joaquin, the observed  $\chi^2_0$ -statistic was found to be [9] 22.2. The value of  $\chi^2$  for significance at 99% level of confidence is 9.2. Therefore, a significant increase was detected by the use of the target-control  $\chi^2$ -test. This shows that for the same set of

data for the target basin, the target-control  $\chi^2$ -test is overwhelmingly more discriminating than the target two-sample t-test and the target two-sample u-test.

2.4 Target-control conditional Student's t-test. In this test population parameters are not known. What is tested is the normality or abnormality of the target, given the behavior of the control, normal or otherwise [9].

Let  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n_1}$  and  $q_{21}$ ,  $q_{22}$ , ...,  $q_{2n_2}$  be the  $n_1$  and  $n_2$  observations of a hydrologic variable in the target watershed before and during seeded periods respectively. Let  $q_{11}^{\prime}$ ,  $q_{12}^{\prime}$ , ...,  $q_{1n_1}^{\prime}$  and  $q_{21}^{\prime}$ ,  $q_{22}^{\prime}$ , ...,  $q_{2n_2}^{\prime}$  be the corresponding observations in the control watershed.

By application of the maximum-likelihood ratio method [25], the test statistic:

$$t_{o} = \frac{\sqrt{n_{1}+n_{2}-3} \left[ (\overline{q}_{2}-\overline{q}_{1}) - (\overline{q}_{2}'-\overline{q}_{1}') \left\{ \sum_{i=1}^{n_{1}} a_{i} (\Delta q_{1i}) + \sum_{i=1}^{n_{2}} b_{i} (\Delta q_{2i}) \right\} \right]}{\left[ \frac{1}{n_{1}} + \frac{1}{n_{2}} + (\overline{q}_{2}'-\overline{q}_{1}')^{2} \right]^{\frac{1}{2}}} \\ t_{o} = \frac{\left[ \sum_{i=1}^{n_{1}} (\Delta q_{1i})^{2} + \sum_{i=1}^{n_{2}} (\Delta q_{2i})^{2} - \left\{ \sum_{i=1}^{n_{1}} a_{i} (\Delta q_{1i}) + \sum_{i=1}^{n_{2}} b_{i} (\Delta q_{2i}) \right\}^{2} \right]^{\frac{1}{2}}}$$

is obtained and it is distributed as Student's t-distribution with  $n_1 + n_2 - 3$  degrees of freedom, where

$$\overline{q}_1 = \frac{1}{n_1} \sum_{i=1}^{n_1} q_{1i}$$

$$\overline{q}_2 = \frac{1}{n_2} \sum_{i=1}^{n_2} q_{2i}$$

$$\overline{q}_{1}^{\star} = \frac{1}{n_{1}} \quad \sum_{i=1}^{n_{1}} q_{1i}^{\star}$$

$$\overline{q}_{2}' = \frac{1}{n_{2}} \sum_{i=1}^{n_{2}} q_{2i}'$$

$$(\Delta q_{1i}) = q_{1i} - \overline{q}_1$$

$$(\Delta q_{2i}) = q_{2i} - \overline{q}_{2}$$

$$(\Delta q_{1i}) = q_{1i} - \overline{q}_{1}$$

$$(\Delta q_{2i}^{!}) = q_{2i}^{!} - \overline{q}_{2}^{!}$$

$$\Delta^{2} = \sum_{i=1}^{n_{1}} (\Delta q_{1i}^{i})^{2} + \sum_{i=1}^{n_{2}} (\Delta q_{2i}^{i})^{2}$$

$$a_{i} = \frac{(\Delta q'_{1i})}{\Delta}$$

$$b_{i} = \frac{(\Delta q'_{2i})}{\Delta} .$$

The use of this test can be found in References [9] and [14].

In Reference [9], the application of the targetcontrol conditional Student's t-test was made for the target, South Fork San Joaquin, and the control, Merced River at Pohono Bridge. The observed  $t_0$ -statistic by this method was 3.80. The value of t for significance at 99% was 2.71. Therefore, a significant increase was the result of this test. Comparison of the results of the above mentioned statistic tests show that the target-control x2-test and the target-control conditional Student's t-test are better tests than the target two-sample t-test and the target sample u-test. Also note that for runoff data from high elevation watersheds the outcomes of the two tests are essentially the same for a sample size around 30. However, it should be noted that all these tests are applicable only when single target or single target-control technique is used. None of these tests can be applied without modification when the number of variables in the study is greater than two, which is the usual case.

2.5 Rank test. Let  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n_1}$  and  $q_{21}$ ,  $q_{22}$ , ...,  $q_{2n_2}$  be  $n_1$  and  $n_2$  observations of a hydrologic variable for the non-seeded and seeded periods respectively.

Arrange the observations in a common sequence of increasing magnitude,  $% \left( 1\right) =\left( 1\right) \left( 1\right) +\left( 1\right) \left( 1\right) \left( 1\right) +\left( 1\right) \left( 1\right) \left$ 

$$q_{11}, q_{12}, q_{21}, q_{22}, q_{13}, q_{14}, q_{15}, q_{23}, q_{16}, \dots$$

Assign ranks from 1 to n , where  $n=n_1+n_2$  , to the above sequence so that rank 1 is given to the smallest observation and n to the largest.

The test statistic is now [26]:

$$Z = \frac{T_s - \overline{T}}{S} ,$$

where Z is approximately a standard normal variate,  $T_s$  is the sum of ranks for seeded observations,  $\overline{T}$  is the expected mean value of  $T_s$ , given by

$$\overline{T} = \frac{n_2(n_2 + n_1 + 1)}{2}$$

$$= \frac{n_2(n + 1)}{2},$$

$$\sigma = \sqrt{\frac{n_2n_1(n + 1)}{12}}$$

If Z is greater than 1.645, then, one rejects the null hypothesis and concludes that at the 5% level of significance weather modification was effective.

This test has been used in References [27] and [28]. From the data in the Necaxa Watershed, Mexico,

it was found that [27] the value of Z was 2.64, which is a value significant beyond the 99% level. The numbers of observations were 45 seeded days and 29 unseeded days. However, the apparent increase in the mean of the seeded period here was large. The seeded mean was about 26 percent larger than the unseeded mean. So, the use of rank test in Reference [27] does not tell much about the efficiency of the test at all. In fact, with the amount of increase of this order, one can find with any statistical test that the cloud seeding is effective. For example, when the u-test is applied the approximate number of observations needed to detect the 26 percent increase in the mean is obtained from:

$$N^* = \frac{4\sigma^2}{h^2 \mu^2}$$

where N\* is the approximate number of observations required to detect a certain amount of increase in the mean,

 $\sigma^2$  is the variance of the hydrologic variable for the unseeded period,

μ is the mean of the hydrologic variable for the unseeded period, and

h is the fractional increase in mean.

Upon substituting the values of  $\,\sigma^2$  ,  $\mu$  ,  $h\,$  from the data of Reference [27], it was found that

$$N^* = \frac{4 \times 600.17}{(.26)^2 (88.14)^2} \approx 5.$$

Thus, it is clear that the required number of observations to detect a 26 percent increase in the mean is much smaller than 45 which is the actual number of observations. So, with this large amount of increase any statistical test will always give the positive result.

2.6 Median test. The median of a distribution is that value which divides the distribution halfway, i.e., half the distribution have lower and half have higher values. The median test determines primarily if the medians of the populations from which the samples come are well separated or not.

Let  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n_1}$  and  $q_{21}$ ,  $q_{22}$ , ...,  $q_{2n_2}$  be  $n_1$  and  $n_2$  observations of a hydrologic variable for the non-seeded and seeded periods respectively. Arrange the observations in a common sequence of increasing magnitude, e.g.,

$$q_{11}$$
,  $q_{12}$ ,  $q_{21}$ ,  $q_{22}$ ,  $q_{23}$ ,  $q_{13}$ ,  $q_{14}$ ,  $q_{15}$ ,  $q_{16}$ ,  $q_{24}$ . . . .

If the total number of observations is even, the median is taken to be halfway between the two middle observations. If this total number is odd, the median observation is removed since it does not contribute any information to the question of whether the distribution of that sample has its median above or below the joint sample median. The case then reduces to the even case.

Let the numbers of  $q_{1i}$ 's above and below the median of the common sequence be  $n_{1a}$  and  $n_{1b}$ , and

the numbers of  $\ q_{2j}$ 's above and below the same common sample median be  $\ n_{2a}$  and  $\ n_{2b}$ . Under the null hypothesis that the two samples come from identical distributions, the proportion of each sample lying below any point should be the same.

If the test function [29]

$$\mathsf{M} = (\left| 2\mathsf{n}_{1a} - (\mathsf{n}_{1a} + \mathsf{n}_{1b}) \right| - 1)^2 / \mathsf{n}_1 + (\left| 2\mathsf{n}_{2a} - (\mathsf{n}_{2a} + \mathsf{n}_{2b}) \right| - 1)^2 / \mathsf{n}_2$$

is greater than  $\chi^2_{0.95}$  with one degree of freedom, then, one rejects, at the 95% level, the hypothesis that the samples have the same median.

This test has been used in Reference [20]. The data used in Reference [20] were obtained from an experiment on artificial stimulation of rain in three climatologically similar regions, Delhi, Agra and Jaipur in northwest India. The net increase in rainfall obtained over all three regions was 41.9%. Thus, it was found that there was a highly significant increase in the amount of rainfall. The observations were made from 1957 to 1965 (excluding 1962) in Delhi, from 1960 to 1965 in Agra, and from 1960 to 1963 in Jaipur. There was, however, no observed statistic given in this report.

2.7 The Mann-Whitney U test. Let  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n}$  and  $q_{21}$ ,  $q_{22}$ , ...,  $q_{2n_2}$  be  $n_1$  and  $n_2$  observations of a hydrologic variable for the non-seeded and seeded periods respectively. Arrange the observations in a common sequence of increasing magnitude, e.g.,

$$q_{11}, q_{12}, q_{13}, q_{21}, q_{14}, q_{15}, q_{22}, q_{23}, q_{24}, \dots$$

The statistic U is defined as the number of times a  $\mathbf{q}_{2j}$  precedes a  $\mathbf{q}_{1i}$  . This test was used to test the null hypothesis

- ${\rm H_{0}}$  the  ${\rm q_{1i}}$  and  ${\rm q_{2j}}$  values have the same distribution against the alternative hypothesis,
- ${
  m H_a}$  the location parameter of  ${
  m q_{2j}}$  is larger than the location parameter of  ${
  m q_{1i}}$ , i.e., the bulk of the distribution of  ${
  m q_{2j}}$ 's is to the right of the bulk of the distribution of  ${
  m q_{1i}}$ 's .

If  $H_a$  is true, one expects U to be small. Mann and Whitney [30] computed tables that give probabilities associated with small (lower tail) values of U, and Auble [31] gives tables of critical values of U for significant levels of 0.001, 0.01, 0.025, and 0.05 for a one-sided test. For the one-sided alternative hypothesis that the location parameter of  $q_{2j}$  is smaller than the parameter of  $q_{1i}$ , one computes the statistic U', defined to be the number of times a  $q_{1i}$  precedes a  $q_{2j}$ , and uses Aubles's tables to test  $H_0$ .

The relationship between U and the sum of ranks for seeded observations,  $T_{\rm S}$ , in the rank test can be expressed as (Wine [32]):

$$U = n_1 n_2 + \frac{n_2(n_2 + 1)}{2} - T_s .$$

The U statistic is usually computed by the above equation, since it is tedious to compute from the definition of U when  $n_1$  and  $n_2$  become fairly large.

The test statistic is

$$W = \frac{U - \overline{U}}{\sigma} ,$$

where  $\frac{W}{\overline{U}}$  is approximately a standard normal variate,  $\overline{U}$  is the expected value of U , given by

$$\overline{U} = \frac{n_1 n_2}{2} \quad ,$$

and 
$$\sigma = \sqrt{\frac{n_1 n_2 (n_1 + n_2 + 1)}{12}}$$

If W is greater than 1.65, then the null hypothesis is rejected and one can conclude the location of  $q_{2j}$  is larger than that of  $q_{1i}$ . This test has been used by many authors - [20], [21], [28], [33], and [34].

In Reference [21], the data used were collected from a five-year period experiment (1960 through 1964) in Missouri. On comparing the average rainfall (inches/hour) of the seeded days with that of the non-seeded days, it was found there was, on the average, a decrease of 67.9%. The values of W ranged from smaller than 0.01 to 0.88. Thus, it was concluded that no evidence of increases in precipitation because of cloud seeding was achieved.

2.8 Run test. Let  $q_{11}$ ,  $q_{12}$ , ...,  $q_{1n_1}$  and  $q_{21}$ ,  $q_{22}$ , ...,  $q_{2n_2}$  be  $n_1$  and  $n_2$  observations of a hydrologic variable for the non-seeded and seeded periods respectively.

Arrange the observations in a common sequence of increasing magnitude, e.g.,

$$q_{11}, q_{12}, q_{21}, q_{13}, q_{14}, q_{22}, q_{23}, \dots$$

A run is defined as an unbroken sequence of elements of the same type, i.e., a sequence of  $q_{1i}$ 's or a sequence of  $q_{2j}$ 's . Let the number of runs be denoted by  $\eta$  . If two samples are from the same population, the non-seeded and seeded observations will be well mixed and the number of runs,  $\eta$ , will be large.

The test statistic is now [14]

$$U = \frac{n - \overline{n}}{\sigma} ,$$

where  $\frac{\underline{U}}{\eta}$  is a standard normal variate, is the expected value of  $\eta$  , given by

$$\overline{\eta} = \frac{2n_1n_2}{n_1+n_2}$$

$$\sigma = \sqrt{\frac{2n_1n_2(2n_1n_2-n_1-n_2)}{(n_1+n_2)^2(n_1+n_2-1)}}$$

If U is greater than 1.65, then the null hypothesis is rejected and the alternative hypothesis is accepted. This test has been used in Reference [35].

In Reference [35], the data of the King River at Piedra, California was analyzed. The observations were the annual flows from 1917 to 1954 for the non-seeded period, and 1955 to 1966 for the seeded period There was a decrease of about 3.3% in mean annual

flows for the seeded period. The number of runs,  $\eta$ , was found to be 17,  $\overline{\eta}$  = 19.240, and  $\sigma$  = 2.533. From the above values, U was obtained as -0.88. Therefore, no significant increase in the mean annual flow was concluded.

Of all the tests stated above, it is found that none of them can be applied for testing the increase in runoff means when the number of runoff variables is greater than two. In the evaluation of weather modification effectiveness based on a multiple target-control concept the number of runoff variables involved is large. So, it is necessary to find an approach to detect the increase in means of these runoff variables.

In Chapter III, the principal components, canonical analysis, and the  $\ensuremath{\mathrm{T}}^2\mbox{-statistic}$  are discussed.

#### Chapter III

#### PRINCIPAL, CANONICAL COMPONENTS AND THE T2-STATISTIC

For small scale operations the method of evaluation of a significant change in hydrologic characteristics based on the single target-control concept is adequate. For large regions this procedure would not be very representative. Besides if the test were performed for many pairs of target and control it is not clear how one should treat the ensemble of the outcomes. On the other hand, there is no problem of interpretation when a single test is performed even though the tested statistic may itself be a complicated combination of many observations from many targets and controls. For representativity the station runoff variables should be geographically well distributed over the large area of interest. This results in a selection of a large number of variables that are usually not independent variables. Sometimes the number of variables involved may be so large that any study can hardly be made economically. In fact, this is one of the difficulties in this study since there are three big areas under investigation. It is, therefore, also an object of this study to find a suitable method for reducing the number of variables involved in the analysis.

There are several ways to reduce the number of variables. However, two methods are used here before the statistical test is carried out. One is the principal components analysis, the other the canonical analysis.

5.1 Principal component analysis. The principal components are linear combinations of random variables, which have special properties in terms of variances. Usually, the linear combination with the maximum variance is referred to as the first principal component; the second component is the one that is uncorrelated with the first and has the second largest variance, and so on. The idea of this analysis was discussed thoroughly by Hotelling [36] in 1933.

From the hydrologic point of view, these principal components can be considered as new transformed runoff variables though lacking simple physical meaning. These transformed variables have, in total, the same amount of fluctuation or variation as do the original runoff variables. But the number of the transformed variables can be smaller than that of the original variables. Also these transformed variables are independent while the original variables are not.

A priori what can be expected from the principal components analysis for the purpose of evaluation? Suppose the principal components analysis is carried for all the targets and all the controls. The first principal component for each group will be the most statistically representative single combination of targets and controls, respectively, because that combination will account for the largest fraction of the total variation. If the percentage is high (say 95%) all the other principal components can be dropped. Then the originally multivariate test reduces again to a familiar single target control t-test, even

though the target variable and the control variable are each a combination of many target and control ones. The procedure will be simple and effective if the target first principal component and the control one are highly correlated. However, this need not happen because the targets and controls are treated separately and the procedure does not attempt to maximize the correlation between the two components (which canonical analysis does). It can be concluded that principal components analysis can provide the basis for a simple and highly representative test but it will not be, by far, a minimal time evaluation one. (The procedure for the actual computation of the principal components is summarized in Chapter V, Section 1).

3.2 The canonical analysis. Canonical analysis is a technique to maximize the correlations between two groups of random variables. This analysis gives new sets of transformed variables as linear combinations of the original runoff variables. The first linear combination of each group will have the highest correlation, and each is uncorrelated with the other linear combinations in its group. The second linear combinations will have the second highest correlation, the third linear combinations will have the third highest correlation and so on. These linear combinations are referred to as canonical variables or components.

In this study the first group is the group of runoff stations in the target region and the second group is made of stations in the control region. This analysis is particularly advantageous for evaluation purposes. The canonical analysis yields a smaller number of variables for the final test, and most importantly it also guarantees high correlations between the variables of the target and control regions.

3.3 Computation of canonical variables. The steps for computing the canonical variables are now described:

Step 1) Compute the covariance matrix,  $\hat{\underline{V}}$ , of the runoff variables of the two sets (target and control). For  $p_1$  runoff stations in the target region and  $p_2$  in the control region, then

$$\frac{\hat{\mathbf{v}}}{\mathbf{v}} = \begin{bmatrix}
\sigma_{11} & \sigma_{12} & \cdots & \sigma_{1p_{1}} & \sigma_{1}(p_{1}+1) & \cdots & \cdots & \sigma_{1}(p_{1}+p_{2}) \\
\sigma_{21} & \sigma_{22} & \cdots & \sigma_{2p_{1}} & \sigma_{2}(p_{1}+1) & \cdots & \sigma_{2}(p_{1}+p_{2}) \\
\vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\
\sigma_{p_{1}} & \sigma_{p_{1}}^{2} & \sigma_{p_{1}}^{2} & \sigma_{p_{1}}^{2} & \sigma_{p_{1}}^{2}(p_{1}+1) & \cdots & \sigma_{p_{1}}^{2}(p_{1}+p_{2}) \\
\sigma_{(p_{1}+1)1} & \sigma_{(p_{1}+1)2} & \cdots & \sigma_{(p_{1}+1)p_{1}} & \sigma_{(p_{1}+1)(p_{1}+1)} & \cdots & \sigma_{(p_{1}+1)(p_{1}+p_{2})} \\
\vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\
\vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\
\sigma_{(p_{1}+p_{2})1} & \sigma_{(p_{1}+p_{2})2} & \cdots & \sigma_{(p_{1}+p_{2})p_{1}}^{2} & \sigma_{(p_{1}+p_{2})(p_{1}+1)} & \cdots & \sigma_{(p_{1}+p_{2})(p_{1}+p_{2})}
\end{bmatrix} (2)$$

The subscripts of  $\sigma$  are the ordering numbers of the stations. The numbers 1 to  $p_1$  are for the  $p_1$  stations in the target region. The numbers  $p_1+1$  to  $p_1+p_2$  are for the  $p_2$  stations in the control region. For example, the subscript 1 will refer to the first station in the target region, while the subscript  $p_1+1$  will refer to the first station in the control region and the subscript  $p_1+2$  the second station in the control region, etc.

 $\boldsymbol{\sigma}_{\mbox{\scriptsize ii}}$  is the variance of the runoff series for station i , defined as,

$$\sigma_{ii} = \frac{1}{N} \sum_{s=1}^{N} (q_{is} - \overline{q}_{i})^{2}$$
, (3)

where N is the number of years of recorded runoff data,  $\begin{array}{c} q_{is} \\ \overline{q}_{i} \end{array} \text{ is the s}^{th} \text{ recorded runoff of station i, and} \\ \vdots \text{ is the mean of the recorded runoff of station i.} \end{array}$ 

 $\boldsymbol{\sigma}_{\mbox{\scriptsize ij}}$  is the covariance of stations i and j , defined as,

$$\sigma_{ij} = \frac{1}{N} \sum_{s=1}^{N} (q_{is} - \overline{q}_i) (q_{js} - \overline{q}_j)$$
 (4)

$$\sigma_{ij} = \sigma_{ji}$$
.

Step 2) Partition the covariance matrix,  $\underline{\hat{V}}$  , such that,

$$\frac{\hat{\mathbf{v}}}{\hat{\mathbf{v}}} = \begin{bmatrix} \frac{\hat{\mathbf{v}}}{11} & \frac{\hat{\mathbf{v}}}{12} \\ \frac{\hat{\mathbf{v}}}{21} & \frac{\hat{\mathbf{v}}}{22} \end{bmatrix} ,$$
(5)

where  $\frac{\hat{V}}{11}$  is a  $p_1 \times p_1$  matrix,

$$\frac{\hat{\mathbf{y}}_{11}}{\hat{\mathbf{y}}_{11}} = \begin{bmatrix}
\sigma_{11} & \sigma_{12} & \dots & \sigma_{1p_{1}} \\
\sigma_{21} & \sigma_{22} & \dots & \sigma_{2p_{1}} \\
\vdots & \vdots & & & \vdots \\
\vdots & \vdots & & & & \vdots \\
\sigma_{p_{1}1} & \sigma_{p_{1}2} & \dots & \sigma_{p_{1}p_{1}}
\end{bmatrix}$$
(6)

$$\hat{\underline{V}}_{12} = \begin{bmatrix}
\sigma_{1}(p_{1}+1) & \sigma_{1}(p_{1}+2) & \cdots & \sigma_{1}(p_{1}+p_{2}) \\
\sigma_{2}(p_{1}+1) & \sigma_{2}(p_{1}+2) & \cdots & \sigma_{2}(p_{1}+p_{2}) \\
\vdots & \vdots & \vdots & \vdots \\
\sigma_{p_{1}}(p_{1}+1) & \sigma_{p_{1}}(p_{1}+2) & \cdots & \sigma_{p_{1}}(p_{1}+p_{2})
\end{bmatrix} (7)$$

$$\underline{\hat{\mathbf{y}}}_{12} = \underline{\hat{\mathbf{y}}}_{21} \tag{8}$$

$$\hat{\underline{V}}_{22} = \begin{bmatrix}
\sigma(p_1+1)(p_1+1) & \sigma(p_1+1)(p_1+2) \cdots \sigma(p_1+1)(p_1+p_2) \\
\sigma(p_1+2)(p_1+1) & \dot{\sigma}(p_1+2)(p_1+2) \cdots \sigma(p_1+2)(p_1+p_2) \\
\vdots & \vdots & \vdots \\
\sigma(p_1+p_2)(p_1+1) & \sigma(p_1+p_2)(p_1+2) \cdots \sigma(p_1+p_2)(p_1+p_2)
\end{bmatrix}$$
(9)

Step 3) Obtain the values of canonical correlations by solving the system,

$$\begin{vmatrix} -\theta \hat{\underline{V}}_{11} & \hat{\underline{V}}_{12} \\ \hat{\underline{V}}_{21} & -\theta \hat{\underline{V}}_{22} \end{vmatrix} = 0 . \tag{10}$$

The values of  $\theta$  are the canonical correlations.

Step 4) Let  $\underline{\alpha}$  and  $\underline{\gamma}$  be the column vectors of coefficients for the canonical variables of the target and control regions respectively. Then, for a given value  $\theta_{\underline{i}}$ , the vectors  $\underline{\alpha}_{\underline{i}}$  and  $\underline{\gamma}_{\underline{i}}$  can be obtained by solving the system,

$$\begin{bmatrix} -\theta_{1} \hat{\underline{V}}_{11} & \hat{\underline{V}}_{12} \\ \hat{\underline{V}}_{21} & -\theta_{1} \hat{\underline{V}}_{22} \end{bmatrix} \begin{bmatrix} \underline{\alpha}_{1} \\ \underline{Y}_{1} \end{bmatrix} = \underline{0}$$
 (11)

subject to the standardization conditions:

$$\underline{\alpha}_{i}^{\prime} \quad \hat{\underline{V}}_{11} \quad \underline{\alpha}_{i} = 1 \tag{12}$$

$$\underline{Y}_{1}^{!} \hat{\underline{V}}_{22} \underline{Y}_{1} = 1 ; \qquad (13)$$

 $\underline{\alpha_i'}$  and  $\underline{\gamma_i'}$  are the transposes of  $\underline{\alpha_i}$  and  $\underline{\gamma_i}$  respectively.

Once the  $\alpha_i$  and  $\gamma_i$  are obtained, the canonical variables for the target region are obtained from the relations:

$$\zeta_i = \underline{\alpha}_i^! \ \underline{Q}_1 \tag{14}$$

where  $\ \varsigma_{i}$  is the  $i^{\mbox{th}}$  canonical variable in the target region

$$\underline{\alpha_{i}} = (\alpha_{i1}, \alpha_{i2}, \dots, \alpha_{ip_{1}})$$
 (15)

$$\underline{Q}_{1} = \begin{bmatrix} Q_{1} \\ Q_{2} \\ \vdots \\ Q_{p_{1}} \end{bmatrix}$$
 (16)

 $\mathbf{Q_1,\,Q_2,\,\ldots,\,Q_{p_1}}$  are runoff variables in the target region.

Similarly,  $\epsilon_{\mbox{\scriptsize i}}$  is the i<sup>th</sup> canonical variable in the control region defined by the relation:

$$\varepsilon_i = \underline{\gamma}_i' \underline{Q}_2$$
 , (17)

where 
$$\underline{Y}_{i}^{!} = (Y_{i(p_{1}+1)}Y_{i(p_{1}+2)}...Y_{i(p_{1}+p_{2})})$$
 (18)

$$\underline{Q}_{2} = \begin{bmatrix}
Q_{p_{1}+1} \\ Q_{p_{1}+2} \\ \vdots \\ Q_{p_{1}+p_{2}}
\end{bmatrix}$$
(19)

 $\mathbf{Q}_{\mathbf{p}_1+1}$  ,  $\mathbf{Q}_{\mathbf{p}_1+2}$  , ...,  $\mathbf{Q}_{\mathbf{p}_1+\mathbf{p}_2}$  are runoff variables in the control region.

an increase in runoff means. In the previous sections two techniques to transform the original runoff varibles were described and in the case of canonical analysis even the basic steps of the procedure were described. However, the multivariate T<sup>2</sup> test applies just as well for the set of original variables. The principal and canonical transformations will either simplify some of the calculations or improve the outcome of the test. Again, the transformations are not necessary to apply the test. Nevertheless in this study the test was only performed for the transformed variables.

Assuming the values of the population mean vector  $\underline{\mu}^{\star}$  and covariance matrix  $\underline{V}$  for the seeded period are known, the minimum number of observations, N\* , that one needs in order to be able to reject the hypothesis  $\underline{\mu}^{\star}$  =  $\underline{\mu}_{o}$  , where  $\underline{\mu}_{o}$  is a given vector, is given by

$$N^* = \frac{\tau^2}{(\underline{\nu}^* - \underline{\nu}_0)! \ \underline{v}^{-1} \ (\underline{\nu}^* - \underline{\nu}_0)} , \qquad (20)$$

where  $\tau^2$  is the noncentrality parameter with degrees of freedom  $\,k\,$  and  $\,N\!-\!k$  ,

k is the total number of runoff variables, and

N is the number of observations for the nonseeded period.

Select values of  $\tau^2$  as given by Tang [37] and Lehmer [38] are shown for convenience in Table 1.

 $\label{eq:table 1 - Value of } TABLE \ 1 \ - \ VALUE \ OF \ \tau^2$  Level of significance,  $\alpha$  = 0.05; power  $\beta$  = 0.50

Degrees o	f freedom	
k	N-k	τ2
2	28	5.468
4	26	7.640
5	25	8.640
6	24	9.646
8	22	11.655

In this study the value of  $\underline{\nu}$  is assumed to be the mean vector of target and control runoff variables for the period before seeding.  $\underline{\nu}^*$  is similar to  $\underline{\nu}$  except that the means of the target runoff variables are 1.1 times greater than those in  $\underline{\nu}$ . In other words, it is assumed in this study that the effect of precipitation management over the target areas will be to increase the runoff uniformly throughout the target areas by 10%. The covariance matrix  $\underline{V}$  is assumed to be the same as that of the nonseeded period.

When the principal components (or the canonical variables) are used for computing N\* , then  $\underline{\nu}^*$  and  $\underline{\nu}$  are the mean vectors of the principal components (or the canonical variables) for the seeded and non-seeded periods respectively, and  $\underline{\nu}$  is the covariance matrix of the principal components (or the canonical variables) for the non-seeded period. The original runoff variables can also be used in computing N\* . However, because of the large number of the original runoff variables, they are not used in this study.

It should be noted here that the use of principal components in equation (20) will yield approximately the same results as the use of the original runoff variables. This is due to the fact that the amount of variation accounted for by the principal components is practically the same as the variation of the original runoff variables. Thus, the principal component analysis will merely reduce the number of original variables, but will not improve the final outcome of the test.

However, if the number of variables can be reduced to one component then the principal component analysis will be very useful because one can apply a bivariate test, such as the conditional Student's t-test which is less restrictive in its assumptions than the  $T^2$ -test. Unfortunately, this usefulness will not be known until one has completed the analysis.

In the next chapter the collection of data in the Upper Basin of the Colorado River, the San Juan Mountains area, and the Maroon Peak and Grand Mesa region is discussed.

#### Chapter IV

#### RESEARCH DATA ASSEMBLY

The data used in this study are the records of the runoff from three regions in the Colorado River Basin. These are:

- 1. The Upper Basin of the Colorado River,
- 2. The San Juan Mountains area,
- 3. The Maroon Peak and Grand Mesa region.

The first two areas were originally [7] proposed as sites for extensive cloud seeding operation. They are called northern and southern target regions (Figs. 2 and 3), while the third is called the control region (Fig. 4). Currently [8] only one area is considered: the southern area. The selection of the control stations is done primarily on the basis of the high correlations with those in the target regions.

It is virgin flow, which is the flow free from any man-made intervention, that is necessary for this study. So, corrections must be made for the records of runoff. The records of runoff were obtained from U.S. Geological Survey Water Supply Papers. However, only the corrections due to transmountain, transbasin diversions, and regulation can be made. The diversion for irrigation cannot be made because there is no record for the amount of water diverted for this purpose. Thus, it is assumed after making the corrections above, that virgin flows are obtained.

Out of a large number of stations, seven stations are chosen for the final analysis in the northern target region, and six stations in the southern region. There are fourteen stations used as controls for the northern region, and nine stations as controls for the southern region. These stations and their descriptions are listed in Table 2. The correlations for these stations computed from all the corresponding actually available records are shown in Tables 3, 4, 5 and 6. There are two stations used as controls for both the northern and southern regions.

Types	Seq.	CSU Sta.	USGS Sta.	TABLE 2 - DESCRIPTION OF STATIONS Names (	0	Lat.	" 1	( n	ong.		Area (Sq. Mi.)	Elevation (ft.)
Types	140.	NO.	40.	, manes				-		-	(34: -11)	(22.7
Target-	1	1970000	9.0105	Colorado River below Baker Gulch, near Grand Lake, Colorado.	54.60	19	33		51	22	53	8750
in	2	1960000	9.0110	Colorado River near Grand Lake, Colo.	40	13	08	105	51	25	103	8380
Northern Project	3	1866000	9.0163	Arapaho Creek at Monarch Lake outlet, Colo.	40	06	45	105	44	57	47.1	8310
A TOWN	4	1830000	9.0190	Colorado River below Lake Granby, Colo.	40	DB	39	105	52	00	311	8050
	5	1820000	9.0195	Colorado River near Granby, Colo.	40	07	15	105	54	00	322	7960
	6	1802730	9.0265	St. Louis Creek near Fraser, Colo.	39	54	30	105	52	45	32.6	8980
	7	1776000	9.0360	Williams Fork near Leal, Colo.	39	49	\$5	106	03	20	89.5	8790
Control-	1	1742100	9.0535	Blue River above Green Mountain	39	49	55	106	13	20	514	7947
stations	250		0.0000	Reservoir, Colo.	20	200	200	107	***	00	590	76.07
for Northern	2	1740000	9.0575	Bluc River below Green Mountain Reservoir, Colo.	39	52	50	106	20	333	33	7683
Project	3	1720000	9.0595	Piney River near State Bridge, Colo.	29	48	00	106	35	00	82.6	7272
	4	1666300	9.0645	Homestake Greek near Red Cliff, Colo.	39	28	25	106	22	00	58.9	8783
	5	1594260	9.0780	Fryingpan River at Norrie, Colo.	28	19	50	106	39	30	89.5	8410
	6	1594236	9.0785	North Fork Fryingpan River near Norrie, Colo.	39	20	40	106	39	50	41.2	8400
	7	1590000	9.0850	Roaring Fork River at Glenwood Springs, Colo.	39	32	50	107	19	50	1460	5721
	8	1379000	9.1090	Taylor River below Taylor Park Reservoir, Colo.	38	45	50	106	36	40	254	9170
	9	1378400	9.1100	Taylor River at Almont, Colo.	38	40	00	106	51	00	477	8011
	10	1378100	9.1125	East River at Almont, Colo.	38	40	00	106	51	00	295	8006
	11	1377825	9.1135	Ohio Creek near Baldwin, Colo.	38	42	00	107	00	00	124	8180
	12	1377500	9.1145	Gunnison River near Gunnison, Colo.	38	32	50	106	57	00	1010	7670
	1.3	1377280	9.1195	Tomichi Creek at Sargents, Colo.	38	24	00	106	25	00	155	8420
	14	1377230	9.1180	Quartz Creek near Ohio City, Colo.	38	33	35	106	38	10	106	8430
			2 1142				-	100		**	100	
Target-	1	1278800	9.1650	Dolores River below Rico, Colo.	37	38	20	108	03	35	105	8422
stations	2	1278050	9.1665	Dolores River at Dolores, Colo.	37	28	00	108	30	00	556	6919
in Southern	3	1272445	9.1725	San Miguel River near Placerville, Colo.	28	02	05	108	07	15	308	7036
Project	4	1077090	9.3440	Navajo River at Banded Peak Ranch, near Chromo, Colo.	37	0.5	07	106	41	20	69.8	7941
	5	1073480	9.3575	Animas River at Howardsville, Colo.	57	50	00	107	36	00	55.9	9617
	6	1073436	9.3615	Animas River at Durango, Colo.		16	45	107	52	47	692	6502
Consuct	- 47	1407627	9.0975	Busined County name Collinson Collinson	39	16	20	107	51	00	139	6955
Control-	1	1425625		Buzzard Creek near Collbran, Colo.								8420
stations	2	1377280	9.1155	Tomichi Creek at Sargents, Colo.	38	24	0.0	106	25	00	155	
for	3	1377230	9.1180	Quartz Creek near Ohio City, Colo.	38	33	35	106	38	10	106	8430
Southern	4	1377200	9.1190	Tomichi Creek at Gunnison, Colo.	38	31	20	106	56	25	1020	7629
Project	5	1373900	9.1275	Crystal Creek near Maher, Colo.	28	22	0.5	107	30	20	42.2	8070
	6	1373055	9,1325	North Fork Gunnison River near Somerset, Colo.	38	55	45	107	26	55	521	6039
	7	1373020	9:1345	Loroux Creek near Cedaredge, Colo.	38	55	35	107	47	35	35.1	7160
	8	1371815	9.1430	Surface Creek near Cedaredge, Colo.	3.5		00	107	51	00	26.7	8180
		1370300		Kannah Creek near Whitewater, Colo.	38		0.0	108	14	0.0	61.9	

TABLE 3 - CORRELATION	MATRIX BETWEEN	N-4 AND CN-4	(as computed	from all	available data)

					N	1-4			
		CSU STA. No.	1970000	1960000	1866000	1830000	1820000	1802730	1776000
	CSU STA. No.	USGS STA. No.	9.0105	9.0105	9.0110	9.0165	9.0190	9.0265	9.0360
CH-4	1742100 1740000 1720000 1666300 1594260 1594236 159000 1378400 1378400 1377500 1377280 1377280	9.0535 9.0575 9.0595 9.0645 9.0780 9.0785 9.0785 9.0850 9.1100 9.1125 9.1145 9.1155	.8625 .6055 .9164 .6781 .8952 .8608 .8723 .8364 .8474 .8635 .8741 .8714 .8026	. 6365 .7277 .9003 .7548 .8514 .8567 .8776 .8174 .8434 .8434 .8434 .8434 .8434 .8434	.8375 .6970 .8322 .8147 .8187 .9187 .8382 .7646 .7971 .5844 .7996 .7937	. 8234 .7077 .8476 .8304 .8919 .9089 .8770 .8541 .8473 .3301 .7306 .8434 .8009 .7675	.6475 .4634 .7171 .6023 .7218 .7975 .7701 .6999 .7329 .6581 .5222 .6851 .6634 .6274	.7779 .8427 .6076 .6515 .6618 .6291 .6381 .4699 .5012 .6456 .6190 .5337 .5672	-9342 -93470 -9470 -8033 -9219 -8647 -8717 -9080 -7744 -7672 -8012 -7082

TABLE 4 - CORRELATION MATRIX BETWEEN N-6 AND CN-6 (as computed from all available data)

						1-6			
		CSU STA. No.	1970000	1960000	1866000	1830000	1920000	1802730	1776000
	CSU STA. No.	USGS STA. No.	9.0105	9.0110	9.0165	9.0190	9.0195	9.0265	9.0360
CN-6	1742100 1740000 1720000 1666300 1694260 1594236 1590000 1378400 1378100 1377825 13777500 1377280 1377280	9.0535 9.0575 9.0595 9.0785 9.0785 9.0785 9.0190 9.1100 9.1125 9.1135 9.1145 9.1180	-6648 -7233 -6944 -7348 -4923 -7153 -7877 -7135 -3529 -7503 -7354 -7354 -8004	.9578 .4359 .2230 .6093 .5923 .5923 .5072 .4373 .4010 .2784 .4247 .5034 .5576	.9124 .8794 .5386 .5738 .6371 .8017 .6076 .6912 .5538 .7701 .6396 .5132 .6163 .1133	.9243 .8640 .6146 .5336 .7567 .6039 .6012 .2488 .7051 .1470 .7202 .5766 .5225 .6640	. 9937 . 9038 . 3806 . 3514 . 8406 . 2842 . 7616 . 4766 . 3055 . 4582 . 0136 . 5478 . 4383	.5921 .4789 .3611 .3718 .5299 .3247 .5203 .5284 .5282 .4277 .7097 .4159 .4590 .6961	-9155 -8409 -3126 -6702 -7359 -4008 -5695 -7748 -7908 -4510 -4439 -8483 -8899 -4868

TABLE S - CORRELATION MATRIX SETWEEN S-4 AND CS-4 (as computed from all available data)

					8	-4		
		CSU STA. NO.	1278800	1278050	1272445	1077090	1073480	1073436
	CSU STA. NO.	USGS STA. NO.	9.1650	9.1665	9.1725	9.3440	9.3440	9.3615
CS-4	1425625 1377280 1377230 1377200 1373900 1373055 1373020 1371815 1370300	9.0975 9.1155 9.1180 9.1190 9.1275 9.1325 9.1345 9.1430 9.1520	.9004 .9020 .9108 .9865 .8879 .8900 .8335	.8519 .7565 .7289 .8587 .8710 .8599 .8608 .8993	.8872 .8040 .5841 .8428 .9059 .7981 .7064 .8021	.7978 .7529 .6336 .7859 .7988 .7835 .8226 .8490	.8466 .8295 .7553 .8895 .8578 .8582 .8118 .2168	.8258 .7353 .6964 .8423 .8549 .8216 .8069 .4315

TABLE 6 - CORRELATION MATRIX BETWEEN S-6"AND CS-6 (as computed from all available data)

			9-6									
		CSU STA. NO.	1278800	1278050	1272445	1077090	1073480	1073436				
	CSU STA. NO.	USGS STA. NO.	9.1650	9.1665	9.1725	9.3440	9.3575	9.3615				
CS-6	1425625 1377280 1377230 1377200 1373900 1373055 1373020 1371815 1370300	9.0975 9.1155 9.1180 9.1190 9.1275 9.1325 9.1345 9.1430 9.1520	.8217 .9310 .8008 .8864 .9406 .9368 .8947 .8844	.6427 .9100 .8536 .7361 .7605 .7148 .6922 .8071	.7111 .7033 .7039 .8601 .8115 .8990 .8129 .7831	.6302 .3573 .7921 .9381 .7576 .8423 .5556 .7217	.8128 .9009 .7754 .9729 .7964 .8881 .8410 .7546	.3267 .9126 .8297 .7719 .7121 .9498 .6934 .7877				

The major part of the spring runoff will occur because of the melting of the winter snow, which is subject to the effect of seeding during winter time. So, it is reasonable to consider whatever changes in the value of the spring runoff as an indirect indicator of the effect of cloud seeding. This is equivalent to saying a larger amount of snowfall in winter will produce a larger amount of runoff in spring. Because of the uncertainty of the start of snow melting, both the runoff during the four months of April, May, June and July, and during the six months of March, April, May, June, July and August are used. These four-month runoff and six-month runoff periods are treated separately in this analysis.

The number of years of record for all stations is fixed at 30, starting from 1938 up to 1967. To assure that these stations are still in operation, the selection has been made in such a way that only stations that have records available for 1967 are considered. It is not likely that the operation of these stations will be discontinued in the near future.

The characteristics of the data used in this study are shown in Tables 7, 8, 9, 10, 11, 12, 13, and 14. There are some data missing in the runoff record of the stations selected but they are filled in by the regression method [39] with the random component superimposed. These stations with missing data are shown in Table 15. Also shown in Table 15 are the stations used in evaluating the missing data. Graphical representations of the data used are shown in Figs. 5, 6, 7, 8, 9, 10, 11, and 12 according to the regions. The means and standard deviations computed from the year 1938 up to 1967 data are shown in Table 16; and the correlations between N-4 and CN-4, N-6 and CN-6, S-4 and CS-4, and S-6 and CS-6 are shown in Tables 17, 18, 19, and 20, respectively.

In Chapter V, the analysis of the data and the results are presented.

			Static	n Numbers			
CSU USGS	1970000 9.0101	1960000 9.0110	1866000 9,0165	9,0190	9,0195	1802730 9,0265	1776000 9.0360
Year							
1938	180,35	300,27	244.61	1175.06	1136.39	102.10	379,77
1939	216.18	245.75	129,61	754.37	765, 51	72.41	247.57
1940	139,53	193.02	168.37	680.57	673.55	53.67	175.50
1941	163.66	251.27	318.69	994.23	805.08	70.50	224.29
1942	216.65	219.05	233,59	1192.95	850.85	83.66	282.53
1943	180,00	230.67	145.06	585,11	807.90	93.99	243.73
1944	191.33	224.75	170.59	661,96	726, 79	73,77	218.40
1945	217,40	251,49	196.92	1060,87	837.71	72,80	233,42
1946	195.31	179,05	201.58	662,49	641.57	73.00	226.37
1947	200,72	317.40	246.16	1556.30	1035.63	98.51	317,49
1948	57.75	193.77	161.58	1023.12	682.34	74.92	231.42
1949	302.67	306.21	244,28	1069.03	1037.22	84,06	268,34
1950	170.28	180.74	155,98	923.67	246_07	72.07	227.31
1951	252.28	305,83	245.06	1000.46	992.64	107.37	301.63
1952	244.15	352.67	285.75	1188.94	1188_71	107.45	334.8
1953	235.87	207.48	186,89	740.89	737.48	75.53	223.91
1954	208.66	117.48	129,00	\$29.31	486.97	34.57	108.27
1955	145.96	150.63	171.00	749.74	438.18	52,04	163.23
195a	206.35	249.24	215.13	1046.18	874.92	42,58	219.86
1957	274.13	364,69	320,64	1625.99	1252.69	84,12	306.45
1958	226.14	249.54	220.52	929.26	995.11	80.52	247.5
1939	171.07	206.87	204.50	453.75	578_10	25.05	201.93
1960	212.75	280,40	217.77	1122.92	1384.95	43.85	252.15
1961	181,57	202,94	171.29	836,40	823.74	23.22	178.7
1962	265, 49	373.30	225,44	1211.50	1105.41	48.1m	313.2
1965	150,16	151.40	158.76	655.72	636.17	21.13	100.4
1964	195, 19	198.73	177.40	826.43	810.53	22.45	177.6
1965	271,56	311.11	264,29	1193.47	1180.51	39.82	275.5
1966	141.71	125.82	99.35	288.08	272.07	20,50	142.6
1967	235.42	243.04	193.79	798.78	788.61	36.67	211.2

			Statio	m Numbers			
CSU	1970000	1960000	1866000	1830000	1820000	1802730	1776000
USGS	9,0105	9,0110	9,0165	9.0196	9.0195	9.0265	9.0360
Year							
1934	126.75	258,12	171.72	806.46	857.32	75.53	269.28
1939	152.42	172.51	91.96	499,87	539,45	53.70	176.43
1940	97.59	135.77	121.40	456.37	479.85	40.21	126,53
1941	114.87	176.89	228.05	683.25	571,12	53,16	162.34
1942	152.78	153.63	166,17	837,27	397.26	62.61	202.05
1943	127.65	164.24	104.99	379.79	578.21	69.92	180.20
1944	133.58	155.78	121.51	426.60	\$11.21	55.45	157.46
1945	159.34	188.16	149.60	773,27	637,78	61,19	183.51
1946	139.16	128.63	143.48	447.69	466.07	55,12	164,47
1947	143.83	228.14	176.76	1116.29	752,17	76,19	237.03
1948	39.11	140,56	114.31	714.13	482.15	55,95	166.26
1949	215.43	216.55	172.61	735.79	731.75	63.60	195.86
1950	119.74	127.12	109.07	638,61	178.45	\$3.53	163,16
1951	185.36	228.48	177.25	723.03	718,66	80.43	221.28
1952	175,29	253.57	205.26	653.98	854.73	81.65	244.36
1953	171.42	154,66	156,06	543.19	542,88	58.32	169.56
1954	77.05	82.52	91.46	377.96	359.63	26, 67	80.90
1955	105.61	112,60	125,39	\$40,96	319.30	41,40	125.25
1956	144.56	277.21	152.30	736.35	621.67	33,12	159.23
1957	201.49	273.95	254.88	1176.53	904,12	59,48	229.60
1958	157.49	175.79	154,87	644.97	710.91	57.91	177.54
1959	123,44	151,96	148,11	618.79	422,89	29.29	147.68
1960	150.52	199.93	153.26	780.54	995,41	31.58	180.31
1961	130.37	148.48	124,20	599,67	589.04	19.44	133,18
1962	193.05	269,76	164.31	851,17	767.64	35.21	225,39
1963	112.60	119,22	120.65	493.81	477,86	16.44	79.03
1964	141.14	146.24	128.08	569.61	\$75.56	18,43	130,81
1965	196.22	227,93	195.31	760.75	749.85	33,52	209,79
1966	101.45	93.69	74.06	109.61	95,96	17.30	106.26
1967	166.64	175.08	135.65	409.44	399.60	27,09	153.63

							Station Nu	shers.						
USGS USGS	1742100 9.0535	1749800 9.0575	9,0595	1666300 9.0645	1594260 9.0780	9.0785	1590000	9,1090	9.1100	1378100 9.1125	9.1135	9,1145	9,1188	9,1180
Year														
1938	1553,47	1425.54	266,05	201,44	335.55	67,04	3970,41	431.36	787,47	909,49	196,74	1527,85	142.26	147,15
1939	664,33	1141.90	226.68	203.33	354,16	142.06	2339,75	\$57,54	\$44.07	574.20	177,54	1514,84	111.77	92.95
1940	766,92	809.29	2,00	149,42	260,23	106.91	1747,83	213,23	302.13	427,80	102,24	1026.65	55,15	48,23
1941	1025.03	1020.70	184.60	178.60	233.90	145.17	2740.32	882.29	592.46	822.97	252,74	21.26	157.12	129,47
1942	1396.12	1191,44	100.01	197.32	243.19	114,37	3125.42	470.89	764.47	754.52	229.56	3676.62	211,14	193,08
1943	619.19	1254,40	61,42	209.79	306,11	164.67	2943.42	451.96	720.99	814.08	272,66	1167.75	148.61	142,62
1944	807.17	973,33	250.99	199.29	277.45	164.86	2897.56	372.48	587.09	774.26	274.26	1700.92	147.20	90.74
1945	865.43	1050.51	194,62	193.80	239.87	. 139,00	2763.15	333,35	490.27	700.26	220,31	1337,85	112.58	88.16
1946	867,13	983.66	109.00	199.61	342.19	121.38	2476.55	326.24	490.30	632.29	2.57.63	1143,95	60.34	61,08
1947	1322.76	1516.34	246,24	256.92	408.12	139.61	3720.73	501.64	755.42	902.67	259.66	1846.62	147.52	136.91
1948	1076,73	1236,84	218.43	222.78	340.15	130,56	3399.91	\$25.59	802.91	941,12	313.44	2022,60	153,99	128,68
1949	1133.94	1322,54	209.23	246.43	362.99	157.57	3093.04	472.44	758.16	824.38	265.90	1782.76	177.18	139,93
1950	926,02	1124.04	156.12	202.52	277.07	103.50	2489.03	354,75	143.30	729.29	216.72	1350.00	73,22	84,37
1951	1377.00	1558.42	200.49	265.24	363.02	146,72	2777.03	397.45	605.49	781.07	210,68	1436.51	120.06	86.09
1952	1294,34	1949.32	270.03	269.57	392.33	176.25	4112,90	629.16	1056.48	1228.04	381.88	2580.67	223.94	178.77
1953	982.09	1156.54	156.10	207,93	285.07	138.78	2476.33	378.85	609.77	648.70	110.59	1345,06	160.10	143.88
1954	435.33	\$18.35	68.95	108.61	156,27	60.74	1268.37	286,50	415,92	371.21	160,29	712.17	33,35	69.57
1955	595.77	726,05	155.83	157.44	233.58	93,63	1950.25	321.17	471.60	\$38.06	122.31	1027,53	55.45	100.22
1956	1019.09	1226,43	170.76	230.74	283,36	133.35	2384.30	412.39	620.21	696,76	218.34	1402,35	98.93	101,33
1957	1339.02	2586,86	290.43	323.31	477,71	246.11	5117.00	765.61	1308.94	1415.11	554,87	3221.64	278.21	216.46
1956	965.24	1167.64	229.26	200.47	298.62	131.64	2725,42	365.70	668,05	861.16	266.45	1704,05	189.62	141.25
1959	605.55	973.27	169.04	189.39	289,58	109.71	2287.30	290.55	419.22	492.07	144,33	940,66	74.47	93,10
1960	865.45	1029.95	182,68	200.25	285.94	130.85	2406.42	418.06	635.63	585.98	177.17	1278.18	122.16	103.95
1961	599.01	727,52	111.20	139.84	209.11	83.65	1768.37	294,09	436.30	477.43	129.06	915.05	90.56	68.80
1962	1084.42	1346,70	291.43	241.40	374.85	162.43	3686.80	597.60	992.49	1048.02	306.57	2345.79	163.62	148.87
1965	422.89	\$35.25	84,07	151.78	176,74	85,67	1413,14	253.52	355.39	417,41	61,39	791,08	57,07	53.64
1964	577.11	396.76	138,68	171.43	256.81	100,15	2240,17	298.35	473.43	\$63,00	136.51	1108.08	125.24	86.33
1965	1259,69	653,70	261.83	286.15	399,81	173.93	3707,34	598.89	1000.70	918.40	350,96	2574,95	195,54	168.81
1966	427.33	520.84	121,10	110.50	202.57	77.59	1835,54	317.30	453.67	524.86	134.20	1034.07	66,46	74,56
1567	624,37	548.06	162.84	59.72	263.00	107,42	2320.36	381.22	\$62.37	708.26	154.24	1306.58	44,29	76,76

							Station Num	mers						
CSU USGS	9.0333	1740000 9,0575	9,0595	1666300 9.0645	1594260 9.0780	1594236 9.0785	1590000 9.0650	1579000 9.1090	9,1100	1378100 9.1125	1377825 9_1135	1377500 9,1145	1377280 9.1155	9,1180
Year														
1938	1156.66	1054,07	184,23	140.66	239,33	44.61	2879.64	316,79	583,73	653,26	139.63	963,01	103,19	108,20
1939	512.72	830,52	156.10	141.35	254.00	98.28	1609.53	264.92	409.91	416.61	127.55	1112.28	82,40	71.4
1940	\$70.41	604,49	.46	103.25	186.31	74.01	1275.78	164.25	236,45	310.62	74.91	769.56	43,81	39,1
1941	760.62	757,80	127.32	124,66	164.67	99,70	1998,73	283.64	440.50	588.59	176.72	15.10	113.88	96.5
1942	1039,35	874.05	67.08	136.77	171.51	77,86	2255.84	350.01	365.74	539.18	161.07	2665.26	150,62	140.2
1941	492,25	933.16	47, 19	151.81	225.05	116.98	2244.66	547.56	556.40	599.26	200,14	902.68	109.85	108.00
1944	595, 54	711,86	172.50	138,46	198,27	114.46	2098.35	277.19	441.90	555.34	192.79	1248,25	106,73	68.8
1945	726,18	256,90	136.97	145.55	179,62	100.35	2157.63	262.68	397.07	521,22	164.32	1046,98	85.38	77.1
1946	640, 28	735,90	119.43	138.42	245.65	83.88	1831.32	242.16	371.95	456,40	112.69	850,04	48,21	48.9
1947	1007.11	1145.85	171.30	182,06	295.19	97.25	2777.18	373.55	568.91	654,97	187.60	1366.27	110.36	101.7
1948	789.57	894,78	151.99	153.3e	240,23	89.65	2486.47	387.42	591.17	669.06	218.91	1454.97	112.85	95.4
1949	847,36	958,77	146.09	171.20	259.18	108.51	2266.56	349.87	562.02	592.63	187,85	1295.83	130.97	104.7
1950	079.63	824,57	108.06	138,51	194,14	70,00	1810.28	261.24	401.21	517.90	151.55	977,46	55.00	63.4
1951	1051.02	1184.87	140,74	188.25	261.82	102.90	2090.26	300.69	463.77	574,60	153,15	1085.94	91.08	67.3
1952	972,44	1164.23	187,50	192,30	284.64	123,11	3045,42	466.85	778,99	879.92	268.12	1874,68	165.79	151.0
1953	744.67	883,46	111.51	147,93	212,46	99.23	1865.67	287.46	464.65	473.36	82.38	1009,12	117.56	107.2
1954	340,53	406.60	48.77	75.41	113.46	42.56	968,70	209.69	318.18	276.42	127.02	549,98	29.41	55.6
1955	492.35	592.03	94,67	113.92	171.63	86.29	1487.11	249,92	372.09	399.65	91.53	797,95	42,97	76.9
1956	753,20	898.64	118.91	158.60	199.63	91.55	1725.41	305.33	459.60	492,50	152.71	1021,19	72.53	76.0
1957	104n, 28	1233.56	200.67	255.32	354.67	172,16	3612.35	589.23	999.55	1051.31	398.44	2421,41	209,08	162.6
1958	705.37	846.10	156.58	137.47	211.06	89,96	1973,10	275.53	501.10	611.32	186,40	1233.51	135,40	103.2
1959	617.20	743.62	118,89	130.20	203.00	75.55	1993.25	220.54	324, 16	361.54	104,17	713,27	59.03	72.5
1960	647,29	771.17	126.52	139.80	202.69	90.29	1771.79	310.01	475.38	423.15	125.97	944,26	92.17	77.1
1961	489.88	591.31	79.11	98.23	151.39	57.94	1320.01	218.30	336.60	345.45	92.44	684.65	69.14	54.3
1962	808.98	997.67	202.09	169.55	265, 87	122.78	2692,93	437.68	728.61	747,02	235,08	1694,99	118.03	108.3
1963	363,95	439,21	61.61	111.42	136.12	61.60	1121.25	202.44	285.55	307,43	62.05	630,30	44.55	44.9
1964	430,63	354,04	97.06	123.74	192.14	71.12	1994.92	232.22	370,13	415,33	99,95	858,04	91,40	66.9
1965	981,95	605.49	188.31	211.24	298.97	125.11	2797.98	462.23	769,46	684,62	251,77	1925.08	143,61	131.4
196e	332.45	420.32	80.49	\$1.82	149.03	55,27	1390.09	243.24	376.24	383.67	97.05	782.29	53,50	59.9
1967	455.68	467.94	114.39	48.23	188.63	75.57	1736.02	285.94	432.59	519,56	116,15	990.10	36.49	61.7

		TABLE	31 - 5-4 583	RIES (CF5)				12	TABLE	12 + 5-6 SEE	KIES (CPS)		
			Station Med							Station Num			
CSU USGS	9.1650	9.1005	1272445 9,1725	9.3440	9.3575	9,3615	CSU	9.1650	1278050 9.1665	9.1725	9,3440	1073480 8,3575	1073436 8.3615
Year							Year						
1938	376.81	1562,90	575,39	325.62	327.16	2336.76	1938	264.70	1087.32	428,99	233,04	234,72	1679.61
			425.55	182,92	209.50	1098.61	1939	32.79	441.02	323.29	136.78	149.07	832.04
1939	50,94	606.38					1940	64.40	335.57	232.45	126,39	138,39	760.03
1940	84.90	745.99	298.12	164.90	197.06	1028.62	1941	438.72	1294.88	799.02	344,08	247.28	2228.01
1941	632,30	1851.80	1089.38	478. 23	336.10	3077.72	1942	228,85	1207.55	403.54	245,38	220,95	1618.76
1942	318,67	1715.27	530,48	338, 20	304,86	2230.25	1943	323.33	520,58	351.01	158,26	172,10	1237,58
1945	449.01	1137.46	422,13	213.89	224.83	1595.65	1944	298.37	1142.23	555.73	221.03	222,45	1864,03
1944	424.04	1639.32	759,53	311.61	308.98	2621.76	1945	183.68	826,26	402.87	208.82	171,12	1270.55
1945	540.66	1162.65	509,43	289,-46	232.80	1706.87	1946	85.90	524.96	320.06	97,90	154.90	915,27
1946	110.67	714.89	408.79	127.24	215.96	1217.53	1947	243,36	767.28	410,77	137.72	215.76	1396.48
1947	329.00	1042,36	517.55	176.18	291.08	1815.22	1948	254.48	948.10	530,04	185,66	241.83	1749.94
1948	358,02	1347.34	726,76	260.75	337.26	2426.32	1949	323.65	972.30	469.79	211.36	242,48	1905,66
1949	456.36	1380.07	638,53	290.50	340,17	2677.44	1950						
1950	92,95	793.92	334,92	148, 56	180.39	1206,87		70.94	371.55	265.88	109.55	128,96	BB6.80
1951	205.16	462,49	260.02	127.78	185,28	928.70	1951	145.70	336.17	201.96	56.99	134.88	707.76
1952	503.09	1859,43	697.05	380.26	365.02	2825.00	1952	389.84	1288.06	520.53	272.82	264.68	2010.24
1953	241,09	647.71	397,84	169,55	180,92	1152,80	1953	173,03	479.38	304.64	126,15	131.97	869.32
1954	194.30	494.48	238.19	144, 24	162.09	1055,70	1954	138.63	361.00	204,30	106.44	117.40	788,64
1953	232.12	638.44	349.38	149.37	164.39	1118.88	1933	172,42	485.31	270.27	134,10	121.39	857.79
1956	248.65	105.34	347,71	161, 12	205.02	1146.48	1956	176,90	494,15	258.79	125.56	144.73	853,34
1957	568.05	1769.26	765.01	179.73	363.13	2551,41	1997	422,62	1291.87	587.51	285.37	275.42	1939.02
							1958	349.47	1046.41	628-61	221.47	205.19	1703.05
1958	506.00	1494.09	588.34	310.35	285.13	2401.89	1959	111,50	268,50	225,53	95.02	229.51	671.83
1959	143.14	352.93	282.65	123, 75	170.30	861,30	1960	237.14	723.01	176_56	182,67	575,83	1265.89
1960	337.76	1020.05	516.91	253, 53	248,82	1735.15	1961	189.47	555,96	361.68	141.86	149,11	1024,97
1961	265.36	770.40	462.98	187.66	206.59	1372.06	1962	251, 22	754,19	378.03	176,50	296,84	1348, 13
1962	361.31	1059.16	511,90	246, 10	273.25	1855.92	1963	144,45	405,77	229,42	103.45	122.02	764,80
1965	190.51	522.54	270.80	129, 95	165.65	957,57	1964	172.04	474,59	305.47	114,14	127, 19	777.66
1964	230,50	650.49	198.27	142.47	175.87	1013.54	1965	367.10	1007.67	517.49	234.55	201.37	1865,86
1965	503,80	1387.75	669,28	325.56	323.32	2487,34							
1900	269.79	852.00	364, 29	219, 23	216,47	1419.96	1966	194,47	633.65	277,44	103.20	154.73	1080.99
1967	180.41	510.44	285.76	170.36	165, 16	973.56	1967	143,22	406.01	235.85	140.37	121.39	771.36

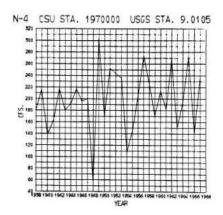
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CSU USGS	1425825 9.0975	9,1155	1377230 9.1186	1377200 9,1190	1573900 9,1275	1373055 9.1325	1573020 9,1345	9,1430	9,1520
lest.									
1938	211.96	142,26	147.15	405.19	114.53	1642,21	205.27	90.56	111,74
1938	74,54	111.77	92,95	262,73	72.35	894,18	100.73	65.04	67.78
1940	65.77	35.15	48.25	89.25	105,32	841.89	98.93	75.51	97.24
1941	182.76	157.12	129.47	410.78	86.48	1395.74	175.09	132.19	148,55
1942	289, 5.3	211.14	193.06	647.95	152,50	1579,42	172.69	123, 79	143.71
1943	73,88	148.61	142.62	363,90	75.68	1059,44	111,46	75. 26	60.63
1944	185,98	147.20	90.74	412,16	120,69	1480.44	139.55	116.49	132,25
1945	130.83	112.58	38.16	261.20	32.12	1249,94	126.28	101.08	87,60
1946	92.40	60.34	61.08	124.30	60.75	847.02	67.73	55.91	56.20
1947	136,23	147,52	136,91	401.64	75.77	1194,08	136.04	113.53	81.28
1946	170,25	153.99	128,68	485,11	115.13	1392,71	118.72	97.64	92.22
1949	154,23	177.18	139.93	582.28	86.42	1239,87	142.84	25.66	96.21
1950	80.95	75,22	84.37	189,60	57.45	1222,23	119.98	84, 06	53.71
1951	42,20	120.06	86.09	239,52	48.28	900.74	83.58	62, 22	44.54
1952	181,21	223.94	178,77	636.87	120.89	1760,47	195.21	166,94	112.23
1953	74.06	160,10	143.88	330.64	77.19	857.61	74.72	64.32	56.70
1954	34.92	33.35	69.57	68.32	22.66	445.60	70.88	57, 12	43.20
1965	59.96	53, 45	100,22	114,93	39,45	619.78	103,11	85, 36	61.01
1956	47,87	98.93	101,35	227.28	51.96	866.61	58.79	61.85	42.93
1057	221.73	278,21	216.46	800.97	147.08	2200.99	259.23	128,75	128.55
1958	185.21	189.82	141.25	496.77	139,64	1297.41	116.86	132.44	125.35
1959	45.48	74.47	93.10	330.29	66.34	660.52	129.36	52, 46	42.65
1960	86.97	122,16	165.95	271.42	45.57	682,67	172.63	71.83	56.4
1961	15.26	90.66	68,80	150.39	46.02	451.57	66.94	69.55	43.7
1962	190.55	163.82	148.87	407,72	95.54	1651.73	161.40	121.98	74.9
1963	23,34	57.07	53.86	107.67	30.48	158.43	66.95	49, 42	19.46
1964	106.35	123,24	86.33	270.43	77.64	989,99	95.81	61.17	49.80
1965	142,90	195,54	168, 61	524.61	112.12	1575.51	131.16	97.77	82.8
1966	62.55	00.46	74,56	156.68	56.62	739.28	85,61	84.05	45,4
1967	41,83	44.79	74.76	164.93	30.83	829.28	105.95	71.72	29.7

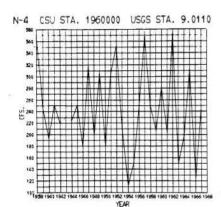
				Stat	ion Numbers				
CSU USGS	1425625 9.0975	9.1155	9.1160	9,1190	1373900 9,1275	1373055 9,1325	1373020 9,1345	137181S 9.1430	9.1520
Year									
1938	144.14	105.19	308,20	302.53	78.66	1136.71	142.02	69.48	80.2
1939	33.13	82.40	71.45	238,90	\$2.07	636.67	70.95	50.37	49.7
1940	46.75	45.61	39.15	91.35	72.57	592.49	69.85	56.37	69.6
1941	123.01	113.66	96.58	376.14	60.40	969.91	122.12	97.97	105.3
1942	194.33	150.62	140.25	474,42	103.98	1090.52	119.28	92.60	101.6
1943	24.77	209.85	108.00	291,91	53.81	758.22	78.60	59.70	45.5
1944	124.48	206.73	58.86	308.07	81.75	1018.14	97.68	69.57	96.0
1945	90,12	85.38	72.17	229.69	24.91	884.45	90.13	80.43	64.5
1946	64.12	48.21	48.99	123, 36	43,47	600.70	49.37	45.81	42.7
1947	92.65	119.36	101.76	320.93	55,40	546.48	96.35	85.87	60.4
1948	115.45	112,35	95.41	365.62	79.24	966.05	84,57	75.55	66.7
1949	90.92	130.97	104,72	435.31	60.49	660.50	101.69	25,39	71.4
1950	55,62	55,06	63,45	152.94	40.33	843.62	80.45	67.46	40.5
1951	29.41	91.08	67.35	196,46	34,44	634.29	61.60	50.05	54.2
1952	122.63	161.78	151.00	468.72	82.59	1220.74	136.39	122.21	80.6
1953	51.39	117.56	107.21	271.75	54.44	613.11	55.97	50, 46	42.7
1954	24,46	29,41	55.67	71.27	15.99	323.00	52.32	45,07	53, 2
1955	40.93	42.97	70.91	109.68	28.76	584.86	74.74	66.55	47.0
1956	33.63	72.53	76.05	174.49	57.67	615.69	64,30	48, 27	33.4
1957	149,42	209.08	162.35	619.51	100.67	1553, 36	181.89	100.35	92.6
1958	126.36	135,40	103.21	365,15	94.29	900.00	82.66	99.83	90.4
1969	31.24	59.03	72.53	123,35	46.51	468.01	91.14	42.37	33.6
1960	60.31	92.17	77,19	219.36	33.19	623, 26	120.63	57,48	49.8
1961	37.06	69.14	54.38	135.32	32.17	460.72	50.14	47.68	30.8
1962	129.10	118.03	108.33	304.76	65,28	1146.48	114,67	65,67	54.3
1965	20.40	48.55	44.90	114.68	25, 22	405.58	48.99	39, 41	15.9
1964	72,26	91,40	66.98	217.55	53.83	693.73	68.87	49.22	35.1
1965	96,84	143.61	131,49	419.56	77.91	1102.29	94.71	76.68	59.7
1966	46,73	53.50	59.98	143.56	41.21	538,26	62.46	65,36	32.4
1967	51.96	36,49	61,74	301,06	23.33	913.85	76.43	58.01	21.5

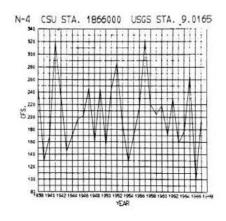
TABLE 15 - STATIONS WITH MISSING DATA

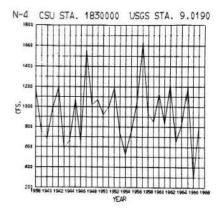
	ion with ing data	Filling in data is mad station		Year of missing data
CSU Sta. No.	USGS Sta. No.	CSU Sta. No.	USGS Sta.	
1820000	9.0195	1830000	9.0190	54-60
1830000	9.0190	1960000	9.0110	38-50
1970000	9.0105	1960000	9.0110	38-53
1272445	9.1725	1277200	9,1665	38-42
1278800	9.1650	1277200	9.1665	38-51
1371815	9.1430	1370300	9.1520	38-39
1373020	9.1345	1373055	9.1325	57-60
1373900	9.1275	1373360	9.1285	38-45; 55-60
1377230	9.1180	1377280	9.1155	51-60
1377825	9.1135	1378100	9.1125	38-40; 51-58
1594236	9.0785	1378400	9.1100	38-47
1720000	9.0595	1590000	9.0850	38-44
1377500	9.1145	1378400	9.1100	38-44
1379000	9.1090	1378400	9.1100	38
1594260	9.0780	1378400	9.1100	38-47

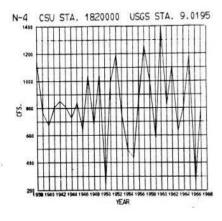
Fig. 5 N-4 series

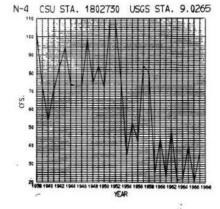












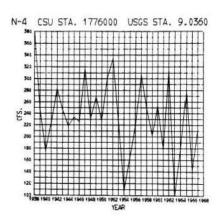
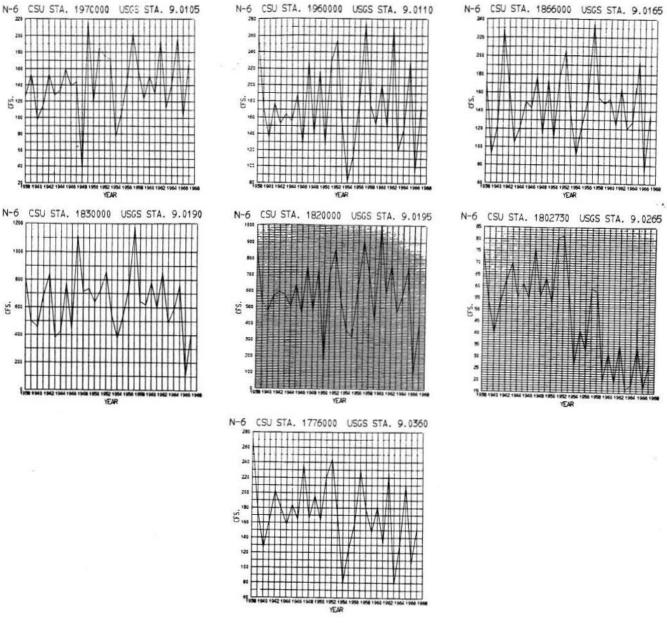


Fig. 6 N-6 series



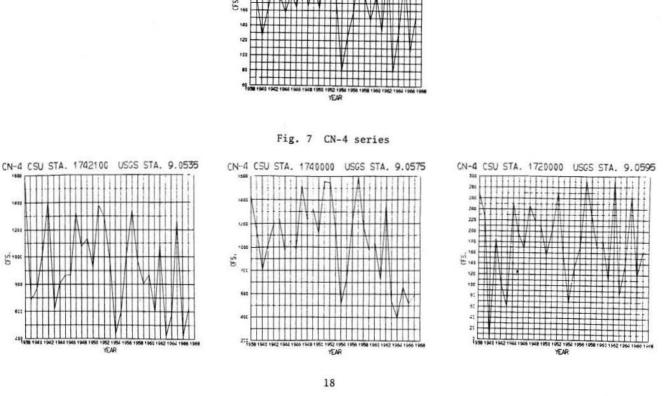


Fig. 7 CN-4 series - Continued

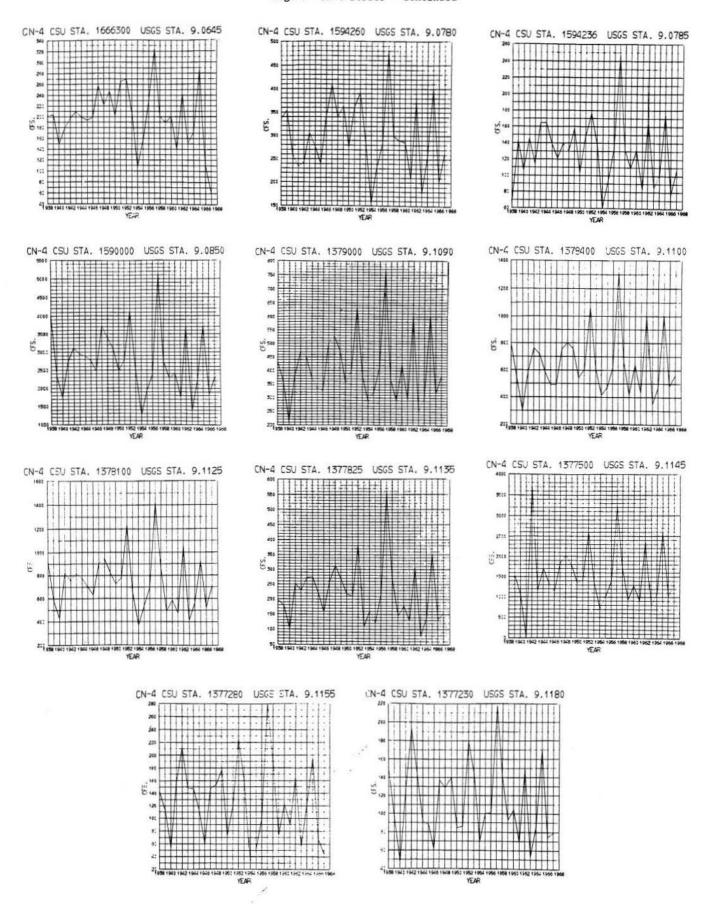


Fig. 8 CN-6 series

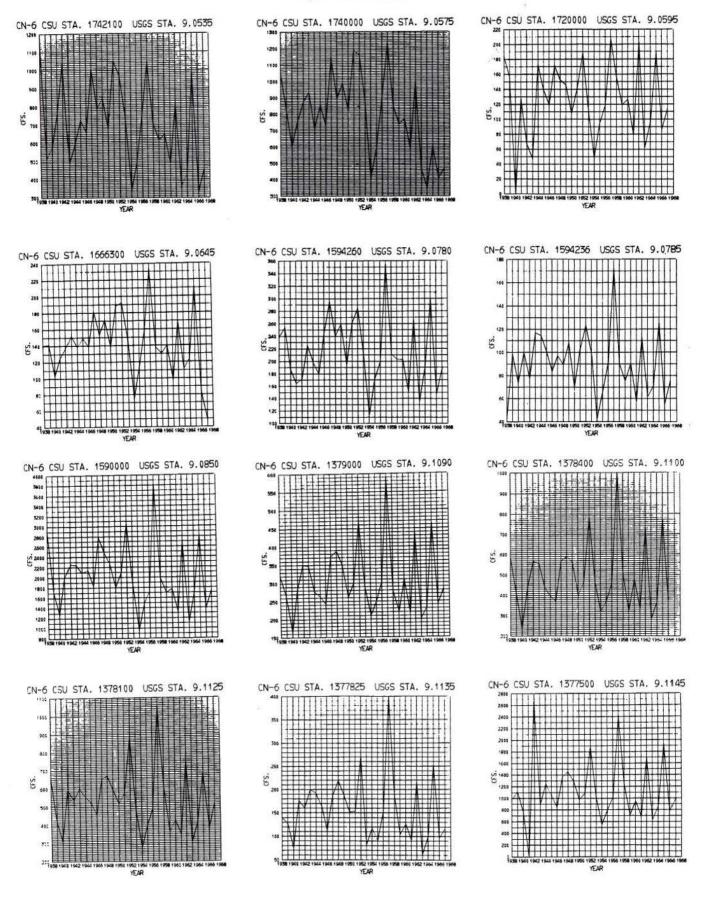
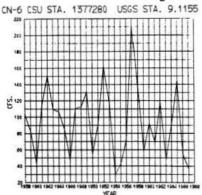


Fig. 8 CN-6 series - Continued



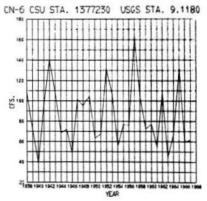
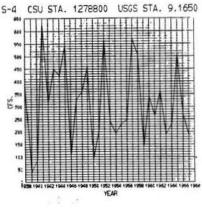
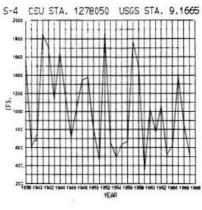
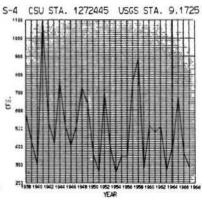
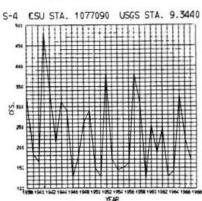


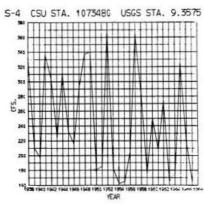
Fig. 9 S-4 series











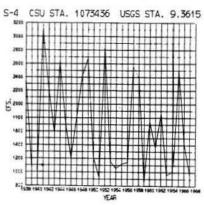
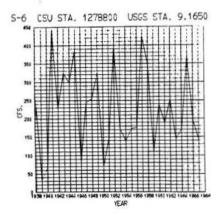
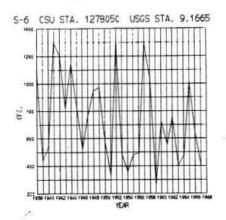
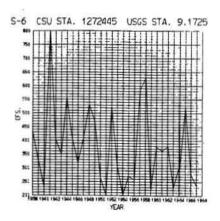


Fig. 10 S-6 series







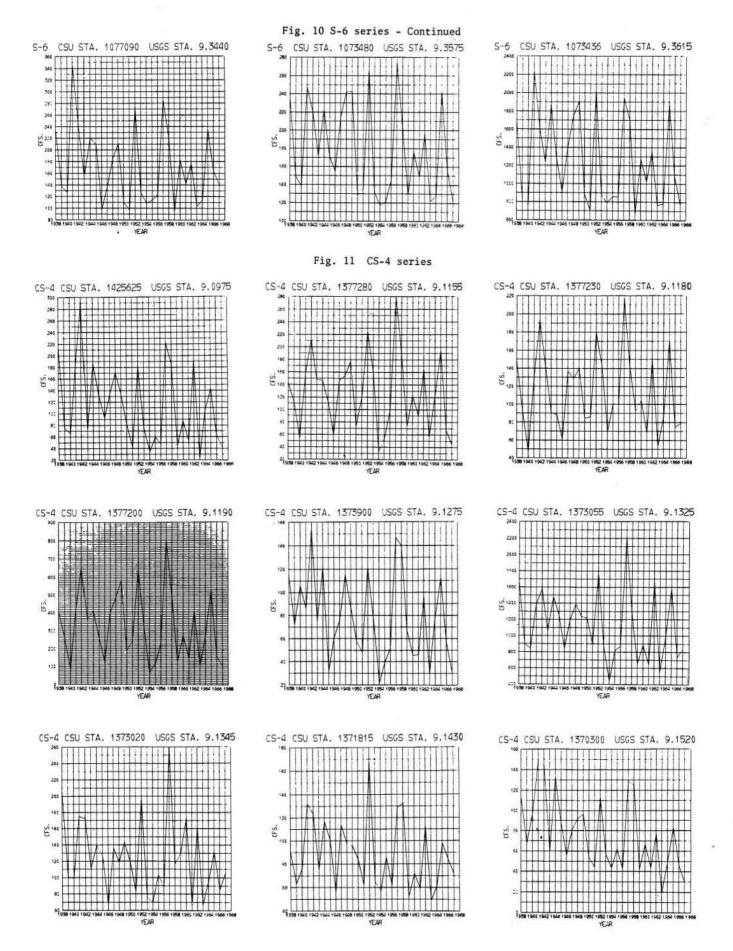
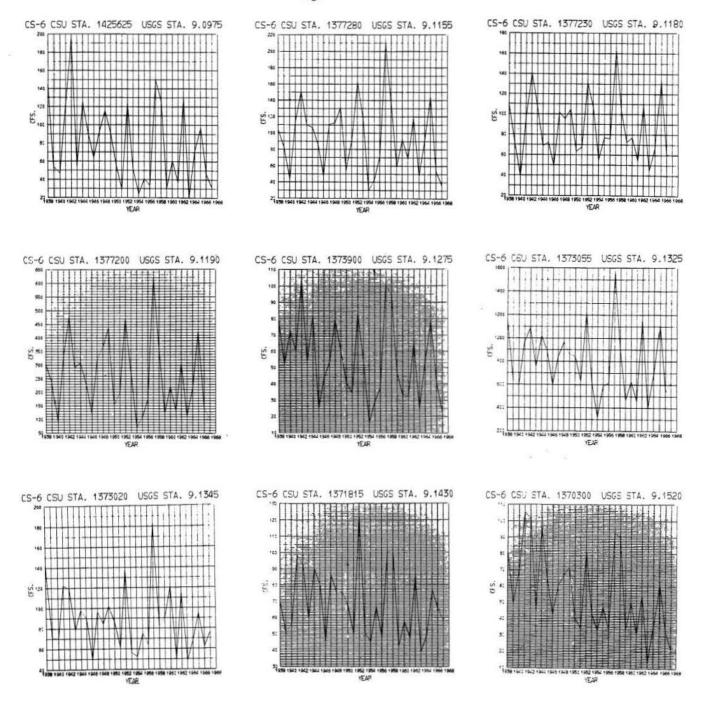


Fig. 12 CS-6 series



Sta. No.	USGS Sta. No.	Mean of 4-month averages (cfs)	Sta. Dev. of 4- month	Mean of 6-month aver- ages (cfs)	Std. Dev. of 6-month aver-
			aver- ages (cfs)	(CLS)	(cfs)
1970000	9.0105	198.449	55.552	141.865	38.338
1960000	9.0110	241.821	71.196	174.569	51.065
1866000	9.0165	203.590	53.612	146.297	38.545
1830000 -	9.0190	931.757	290.050	644.206	220.359
1820000	9.0195	826.556	274.385	582.724	203.842
1802730	9.0265	63.007	27.307	47.814	20.085
1776000	9.0360	234.679	64.188	171.954	45.884
1742100	9.0535	924.237	316.280	702.783	233.361
1740000	9.0575	1043.263	341.708	794.292	241.819
1720000	9.0595	177.674	72.281	124.358	50:072
1666300	9.0645	199.184	55.142	141.030	39.140
1594260	9.0780	297.711	73.851	215.045	53.847
1594236	9.0785	128.576	38.515	89.611	27.138
1590000	9.0850	2739.444	854.102	2031.847	624.926
137900	9.1090	406.685	122.728	306.039	91.987
1378400	9.1100	641.932	224.886	486.119	165.435
1378100	9.1125	736.162	236.020	534.129	170.870
1377825	9.1135	219.301	98.861	157.011	69.485
1377500	9.1145	1521.405	754.366	1129.644	546.188
1377280	9.1155	126.388	59.684	94.178	42.221
1377230	9.1180	113.268	42.698	85.892	30.376
1278800	9.1650	314.930	159.674	226.413	111.344
1278050	9.1665	1028.025	467.197	738.323	319.454
1272445	9-1725	500.048	204.184	379.190	144.622
1077090	9.3440	230.964	93.315	170.480	64.553
1073480	9-3575	245.702	68.310	178.304	50.207
1073436	9.3615	1696.563	688.607	1254.713	481.412
1425625	9.0975	114.324	68.267	78.467	45.159
1377280	9.1155	126.388	59.684	94.178	42.221
1377230	9.1180	113.268	42.698	85.892	30.376
1377200	9.1190	322.737	192.349	257.047	136.341
1373900	9.1275	78.756	36.701	55.056	24.150
1373055	9.1325	1124.250	410.613	790.065	280.861
1373020	9.1345	124.166	46.158	88.247	31.349
1371815	9.1430	88.115	29.509	68.390	21.124
1370300	9.1520	76.695	35.297	56.029	24.863

TABLE 17 - CORRELATION MATRIX SETWEEN N-4 AND CN-4 (computed from 30-year data)

						-4			
		CSU STA. NO.	1970000	1960000	1866000	1830000	1820000	18:-2730	1776000
	CSU STA. NO.	USGS STA. NO.	9.0105	9.0110	9.0165	9.0190	9.0195	9.0265	9.0360
	1742100	9.0535	.477	.785	.778	.815	.641	.710	.894
	1740000	9.0575	.411	.728	-566	.669	.531	.843	.836
	1720000	9.0595	.524	.771	.592	.660	.593	.421	.730
	1666300	9.0645	.535	.717	.640	.728	-619	-592-	.721
	1594260	9.0780	.582	.802	.597	.702	.640	.585	.805
	1594236	9.0785	.621	.652	.602	. 565	.584	.477	.549
CN-4	1590000	9.0850	.502	.845	.722	.806	.673	.641	.870
	1379000	9.1090	-517	.767	.655	.754	.662	.469	-730
	1378400	9.1100	.541	.789	. 671	.743	.685	.489	.757
	1378100	9.1125	.497	.801	.714	.740	.621	-627	.791
	1377025	9.1135	. 426	.680	. 639	.686	.564	.498	.627
	1377500	9.1145	. 497	. 533	.419	.632	. 491	.395	.602
	1377280	9.1155	. 553	.710	.692	.694	. 694	.567	.708
	1377230	9.1180	. 490	.674	.674	.704	.612	.537	.719

TABLE 18 - CORRELATION MATRIX BETWEEN 8-6 AND CN-6 (computed from 30-year data)

						-6			
		CSU STA. NO.	1970000	1960000	1866000	1830000	1820000	1802730	1776000
	CSU STA. NO.	USGS STA. NO.	9.0105	9.0110	9.0165	9.0190	9.0195	9.0265	9.0360
	1742100	9.0535	.486	.783	.780	.813	.671	.706	.896
	1740000	9.0575	-457	.756	.598	.731	.631	.825	.855
- 1	1720000	9.0595	.529	.770	.587	.620	. 569	.408	.736
	1666300	9.0645	-570	.740	.659	.750	.668	.578	.745
	1594260	9.0780	.598	. 798	.592	.659	.621	. 555	. 803
	1594236	9.0785	.627	.650	.603	.529	.552	.448	.562
CN-6	1590000	9.0850	. 509	.840	.720	.764	.659	-623	-876
	1379000	9.1090	. 529	.766	.658	.704	.613	.438	-734
	1378400	9,1100	.551	.788	.669	.693	.639	.456	.761
	1378100	9.1125	.507	.802	.716	.700	.597	.599	. 7:99
	1377825	9.1135	. 433	.681	-640	.650	.541	.477	.644
	1377500	9.1145	.503	.532	.420	. 596	.450	.373	- 6:04
	1377280	9.1155	. 555	.708	.701	.689	.695	.542	.715
	1377230	9.1180	.497	.668	.673	. 674	.593	-509	.7.23

TABLE 19 - CORRELATION MATRIX NETWEEN 5-4 AND CS-4 (computed from 30-year data

		- 7			S-4			
		CSU STA. NO.	1278800	1278050	1272445	1077090	1073480	1073436
	CSU STA. NO.	USGS STA. NO.	9,1650	9.1665	9.1725	9.3440	9.3575	9.3615
	1425625	9.0975	.656	.890	.752	.811	.849	.830
	1377280	9.1155	.748	.807	.718	.753	.830	.782
	1377230	9.1180	.682	.742	. 593	.695	. 760	.709
	1377200	9.1190	.748	.859	.737	.786	.889	.842
CS-4	1373900	9.1275	.491	.776	.661	.663	.770	-714
10000	1373055	9.1325	.707	.961	.711	.792	.858	.810
	1373020	9.1345	.643	.776	.612	.773	.777	.723
	1371815	9.1430	.772	. 894	.784	.836	. 819	.850
	1370300	9.1520	.675	.908	.835	.866	. 828	.856

TABLE 20 - CORRELATION MATRIX RETWEEN S-6 AND CS-6 (computed from 30-year data)

					5-6			
		CSU STA. MO.	1278800	1278050	1272445	1077090	1073480	1073436
	CSU STA. NO.	USGS STA. NO.	9.1650	9.1665	9.1725	9.3440	9.3575	9.3615
CS-6	1425625 1377280 1377230 1377200 1373900 1373900 1373055 1373020 1371815	9.0975 9.1155 9.1180 9.1190 9.1275 9.1325 9.1345 9.1430 9.1520	.653 .753 .702 .753 .492 .716 .656 .777	.892 .805 .747 .854 .778 .863 .779 .888	.754 .722 .608 .746 .659 .720 .624 .793	.804 .747 .699 .783 .660 .791 .780 .837	.846 .840 .779 .894 .770 .866 .787 .828	.826 .786 .722 .841 .712 .813 .733 .854

#### Chapter V

#### DATA ANALYSIS AND RESULTS

In this chapter the data described in Chapter IV are analyzed according to the procedures discussed in Chapter III. The approaches used for reducing the number of runoff variables are the principal component analysis and the canonical analysis. The minimum numbers of years to detect the increase in the runoff means are obtained by application of equation (1).

In the principal component analysis and the canonical analysis, the coefficients for the principal components and the canonical variables are obtained basically from the analysis of the covariance matrix. Therefore, because the covariance matrix is assumed to be the same for both periods, it follows that the coefficients obtained for the non-seeded period apply for the seeded period as well. The suspected change in the means of the runoff leave the coefficients of the components invariant.

5.1 The application of principal component analysis. The numerical procedures for the reduction of the number of runoff variables by the principal components method were executed separately in each region on the CDC 6400 digital computer of Colorado State University. The program BMDOIM from the University of California Press was modified to accommodate nonstandardized variables. The zero mean is not desirable here because a certain percent increase in the mean will be postulated later.

The steps in obtaining the principal components in each region may be summarized as follows:

- 1) Compute the covariance matrix of the runoff variables in that region,  $\hat{\underline{V}}$  , as defined in equation (2).
  - 2) Solve the system,

$$|\hat{\mathbf{V}} - \lambda \mathbf{I}| = 0 \quad , \tag{21}$$

to obtain  $\lambda_1$ ,  $\lambda_2$ , ...,  $\lambda_p$  , the characteristic roots, which are the amounts of variances of components 1, 2, ..., p .

3) Solve the system,

$$(\hat{\underline{V}} - \lambda_i \ \underline{I})\underline{\beta}_i = \underline{0} \tag{22}$$

subject to the normalization condition,

$$\underline{\beta}_{i}^{!} \underline{\beta}_{i} = 1 \tag{23}$$

to obtain  $\underline{\beta}_i$  which is the vector of the coefficients for the i<sup>th</sup> component in that region.

For example, when N-4, which is the four-month runoff of the northern region, is used the coefficients for the first principal component are found to be (Table 21),

$$\beta_{1,1} = 0.0859$$
 $\beta_{1,2} = 0.1679$ 
 $\beta_{1,3} = 0.1151$ 
 $\beta_{1,4} = 0.7065$ 
 $\beta_{1,5} = 0.6576$ 
 $\beta_{1,6} = 0.0332$ 
 $\beta_{1,7} = 0.1359$ 

where the first subscript of  $\beta$  indicates the ordering number of the principal component, the second one indicates the sequential number of the station as shown in Table 2.

Let  $\xi_i$  be the i<sup>th</sup> principal component in the target region before seeding, then for N-4,

$$\xi_1 = \sum_{j=1}^{7} \beta_{1,j} Q_j$$

$$= 0.0859Q_1 + 0.1679Q_2 + 0.1151Q_3 + 0.7065Q_4$$

$$+ 0.6576Q_5 + 0.0332Q_6 + 0.1359Q_7$$

where  $\mathbf{Q_1}$ ,  $\mathbf{Q_2}$ ,  $\mathbf{Q_3}$ ,  $\mathbf{Q_4}$ ,  $\mathbf{Q_5}$ ,  $\mathbf{Q_6}$  and  $\mathbf{Q_7}$  are runoff variables listed in order corresponding to the numbers in the 'Seq. No.' column in Table 2. This first principal component will account for the largest percentage of the total variation in this whole region based on the four-month spring runoff.

The coefficients for the principal components in N-4, N-6, CN-4, CN-6, S-4, S-6, CS-4 and CS-6 are shown in Tables 21, 22, 23, 24, 25, 26, 27, and 28, respectively. The cumulative percentages of total variation accounted for by the principal components in each region are shown in Table 29. A 99 cumulative percentage was used to limit the number of the principal components to be retained for the study, because it was found that beyond this percentage of total variation, the rate of increase of the cumulative percentage was very slow.

After the coefficients of the principal components in each region have been found, then the series of the principal components can be simply obtained from the original series [12].

TABLE 21 - COEFFICIENTS FOR THE PRINCIPAL COMPONENTS OF N-4

CSU	USGS				
Sta.	Sta.	lst	2nd	3rd	4th
No.	No.	Comp.	Comp.	Comp.	Comp.
1970000	9.0105	.0859	0894	4339	8081
1960000	9.0110	.1679	0529	4719	.0637
1866000	9.0165	.1151	.0334	1221	2757
1830000	9.0190	.7065	.6848	.1407	0308
1820000	9.0195	.6576	7201	.1966	.0688
1802730	9.0265	.0332	.0191	3072	.2822
1776000	9.0360	.1359	.0132	6491	.4262

TABLE 22 - COEFFICIENTS FOR THE PRINCIPAL COMPONENTS OF N-6

Sta. No.	USGS Sta. No.	1st Comp.	2nd Comp.	3rd Comp.	4th Comp.
1970000	9.0105	.0767	0806	5604	6494
1960000	9.0110	.1549	0680	5037	.0808
1866000	9.0165	.1084	.0256	2105	2926
1830000	9.0190	.7191	.6784	.1135	0377
1820000	9.0195	.6510	7266	.2122	.0154
1802730	9.0265	.0339	.0048	1892	.4046
1776000	9.0360	.1279	0079	5424	. 5664

TABLE 23 - COEFFICIENTS FOR THE PRINCIPAL COMPONENTS OF CN-4

CSU Sta. No.	USGS Sta. No.	1st Comp.	2nd Comp.	3rd Comp.	4th Comp.
1742100	9.0535	.2250	1496	4782	7580
1740000	9.0575	.2155	2714	7505	5419
1720000	9.0599	.0444	0508	.0329	0134
1666300	9.0645	.0378	0229	0434	.0232
1594260	9.0780	.0524	0388	0108	.0455
1594236	9.0785	.0240	0114	.0178	.0891
1590000	9.0850	.7025	4654	.3580	0949
1379000	9.1090	.0971	0158	.0912	.1170
1378400	9.1100	.1803	0355	.1781	.1584
1378100	9.1125	.1862	1309	.1338	.2226
1377825	9.1135	.0733	0277	.0926	.1475
1377500	9.1145	.5637	.8144	0985	.0404
1377280	9.1155	.0444	0050	.0128	.0178
1377230	9.1180	.0320	0011	.0114	.0005

TABLE 24 - COEFFICIENTS FOR THE PRINCIPAL COMPONENTS OF CN-6

CSU	USGS				
Sta.	Sta.	lst	2nd	3rd	4th
No.	No.	Comp.	Comp.	Comp.	Comp.
1742100	9.0535	.2268	1513	5634	7013
1740000	9.0575	.2156	2566	6879	.6239
1720000	9.0595	.0422	0478	.0304	0128
166300	9.0645	.0375	0220	0413	.0318
1594260	9.0780	.0528	0362	0035	.0583
1594236	9.0785	.0233	0100	.0220	.0898
1590000	9.0850	.7062	4663	.3442	1245
1379000	9.1090	.1005	0147	.1038	.1050
1378400	9.1100	.1825	0337	.1947	.1356
1378100	9.1125	.1858	1261	.1495	.2038
1377825	9.1135	.0710	0262	.0944	.1331
1377500	9.1145	.5577	.8196	0869	.0420
1377280	9.1155	.0433	0047	.0120	.0192
1377230	9.1180	.0313	0008	.0125	0051

TABLE 25 - COEFFICIENTS FOR THE PRINCIPAL COMPONENTS OF S-4

CSU	USGS			
Sta.	Sta.	lst	2nd	3rd
No.	No.	Comp.	Comp.	Comp.
1278800	9.1650	1608	0738	8889
1278050	9.1665	5304	.8066	0525
1272445	9.1725	2180	4039	2817
1077090	9.3440	1027	.0634	1532
1073480	9.3575	0754	0045	.1153
1073436	9.3615	7931	4205	.3017

TABLE 26 - COEFFICIENTS FOR THE PRINCIPAL COMPONENTS OF S-6

CSU Sta. No.	USGS Sta. No.	1st Comp.	2nd Comp.	3rd Comp.
1278800	9.1650	1622	1421	8496
1278050	9.1665	5207	.8186	1102
1272445	9.1725	2240	3730	3252
1077090	9.3440	1013	.0618	1660
1073480	9.3575	0802	0038	.1155
1073436	9.3615	7973	4084	.3456

TABLE 27 - COEFFICIENTS FOR THE PRINCIPAL COMPONENTS OF CS-4

CSU Sta. No.	USGS Sta. No.	1st Comp.	2nd Comp.
1425625	9.0975	1341	.0453
1377280	9.1155	1167	. 2714
1377230	9.1180	0799	.1717
1377200	9.1190	3879	.8378
1373900	9.1275	0658	.0819
1373055	9.1325	8906	4286
1373020	9.1345	0859	0353
1371815	9.1430	0537	.0138
1370300	9.1520	0616	.0331

TABLE 28 - COEFFICIENTS FOR THE PRINCIPAL COMPONENTS OF CS-6

CSU Sta. No.	USGS Sta. No.	lst Comp.	2nd Comp.
1425625	9.0975	1278	0209
1377280	9.1155	1201	2711
1377230	9.1180	0831	1661
1377200	9.1190	3994	8374
1373900	9.1275	0625	0688
1373055	9.1325	8857	.4363
1373020	9.1345	0853	.0376
1371815	9.1430	0567	+.0003
1370300	9.1520	0621	0283

TABLE 29 - CUMULATIVE PERCENTAGE OF TOTAL VARIATION ACCOUNTED FOR BY THE PRINCIPAL COMPONENTS

7	'ype	Principal component	Cumulative percentage of total variation accounted for
		ε1	85
N-4	series	\$1 and \$2	97
	Derres	$\xi_1, \xi_2$ and $\xi_3$	98
		$\xi_1, \xi_2, \xi_3$ and $\xi_4$	99
		ε <sub>1</sub>	85
N-6	series	\$1 and \$2	97
		\$1.\$2 and \$3	98
		$\xi_1, \xi_2, \xi_3$ and $\xi_4$	99
		n <sub>1</sub>	82
CN-4	series	n <sub>1</sub> and n <sub>2</sub>	94
		n1,n2 and n3	98
		71,72,73 and 74	99
		ηı	83
CN-6	series	n <sub>1</sub> and n <sub>2</sub>	95
		n1, n2 and n3	98
		"1'"2'"3 and "4	99
		ε <sub>1</sub>	97
S-4	series	\$1 and \$2	98
		$\xi_1, \xi_2$ and $\xi_3$	99
		ε <sub>1</sub>	97
S-6	series	\$1 and \$2	98
		$\xi_1, \xi_2$ and $\xi_3$	99
		n <sub>1</sub>	95
CS-4	series	n <sub>1</sub> and n <sub>2</sub>	99
		n <sub>1</sub>	95
CS-6	series	n <sub>1</sub> and n <sub>2</sub>	99

The means and standard deviations of the series of the principal components for N-4, N-6, CN-4, CN-6, S-4, S-6, CS-4 and CS-6 are given in Table 30.

It is simply proven [12] that if all the means in the target areas during the seeded period have been increased by a certain fraction of the old means, say h, that is, the increase of  $\mathbf{Q}_1$  is  $h\mathbf{Q}_1$ , of  $\mathbf{Q}_2$  is  $h\mathbf{Q}_2$ , and so on, then the increase in the means of the principal components will also be h . If h is assigned a value of 0.10, then

$$E\{\xi_{i}^{*}\} = 1.1 \ E\{\xi_{i}\}$$
 ,

where E{} denotes the expected value of {}, which is the cloud seeding effect assumed in this study.

TABLE 30 - MEANS AND STANDARD DEVIATIONS
OF THE PRINCIPAL COMPONENTS

Туре		Principal component	Mean (cfs)	Std. Dev. (cfs)
_		ξ1	1316.896	385.728
N-4	series	£ 2	23.431	144.526
		52	-103.167	50.427
		£ 2 £ 3 £ 4	-55.122	37.622
		£2	919.996	289.498
N-6	series	52	-7.066	106.279
		52	-103.770	39.834
		\$1 \$2 \$3 \$4	-19.400	27.483
		n <sub>1</sub>	3566.570	1171.757
N-4	series	n2	-616.325	446.487
		n <sub>3</sub>	-124.669	238.558
		n4	41.293	146.087
		<b>n</b> 1	2656.142	853.439
CN-6	series	n <sub>2</sub>	-442.328	326.332
		n3	-118.678	169.248
		n <sub>4</sub>	49.830	101.846
		ξ1	-2092.706	865.153
5-4	series	52	-95.873	108.688
		ξ <sub>2</sub> ξ <sub>3</sub>	30.022	81.462
		ξ,	-1538.060	601.913
5-6	series	£ 2	-71.786	74.436
		\$1 \$2 \$3	28.886	55.548
CS-4	series	n <sub>1</sub>	-1190.877	459.172
50.000		n2	-146.711	86.254
CS-6	series	ņı	-849.227	315.594
		n 2	85.939	61.805

For the control region, it is obvious that following the assumption that the means of the runoff stations in the control region remain unchanged,

$$E\{n_i^*\} = E\{n_i\}$$

where  $n_i^*$  is the  $i^{th}$  principal component of the control region during the seeded period.

After the principal components in each separate region have been obtained, they are gathered into four major target-control groups as N-4 and CN-4, N-6 and CN-6, S-4 and CS-4, and S-6 and CS-6. For brevity, after the principal components in the target are combined with those in the control, the following symbols will be used:

N-CN-4 - the combination of N-4 and CN-4

N-CN-6 - the combination of N-6 and CN-6

S-CS-4 - the combination of S-4 and CS-4

S-CS-6 - the combination of S-6 and CS-6.

Since it is the principal components that will be utilized in the final test, the computations of the covariance matrices are carried out for these principal components. These are as shown in Tables 31, 32, 33, and 34; also shown are the correlations matrices in Tables 35, 36, 37, and 38.

Same			P-4		CH-4			
	ι,	42	1,3	4	*1	*2	*3	*
- 61	148706,103	-5.194	+1.661	-4.525	361585.644	-32549.149	-7106.836	-1545.409
B-4 12	-5.154	20887.828	361	975	31294.727	1427.066	-4669.399	-2250.743
۲,	-1.541	361	2542.944	,155	-18240.198	5629.908	2877.776	-479.743
	-4.525	075	.155	1415.475	4402.075	-4938.012	-1612.064	149.454
	361585.644	31294.727	-18240.198	4402.075	1373015.745	-71.56	57.238	-49.204
CB-4 "2	-32549.149	1427,066	5429.908	-4938.613	-71.561	199350.870	.050	-3.725
*	-7104.854	-4164.399	2877.776	-1612.064	57.238	.050	56910.290	2.404
	-8545.409	-2250.743	-679.743	149.454	-49,204	-3.719	2.404	21341.511

				M-6		C9-6			
_		1,	4	£3	14	*1	*2	*3	74
	t,	93809.330	3,376	15.020	1.313	191098.859	-20577.629	-10540.852	-2678.411
		3,376	11195.431	2.644	.066	13712.454	3162.736	-2003.334	-963.146
H-6		15.020	2.544	1586.754	.068	-12182.229	1478.320	24.742	-126.57
	i.	1.313	,044	.068	755.354	3884.894	-2718.278	-993.925	121,930
		191090.059	13712.454	-12102,229	3004.094	724334.925	20,36%	20.167	-14.251
	12	-20577,629	3162,736	1678.320	-2710.278	20.301	106493.191	-5.314	91
CB-6		-10540.852	-2003.334	24,742	-993,925	20,187	-5.314	28645.205	06
		-2678, 418	-963,144	-126.575	121.930	-14,259	917	067	10372.75

TABLE 33 - COVARIANCE MATRIX OF S-CS-4 PRINCIPAL COMPONENT SERIES

			S-4		CS-4		
	- 10	ι,	t <sub>2</sub>	43	"1	<sup>n</sup> 2	
	ζ,	748491.385	-45.282	41.666	338072.405	-12524.935	
S-4	62	-45.282	11813.224	485	-11808.907	-1953.006	
	43	-41.666	485	6636.209	-636.724	-167.057	
	*1	338072.405	-11808.907	-636.724	210839.108	7.238	
C5-4	72	-12524.935	-1953.006	-167.057	7.238	7439.863	

TABLE 34 - COVARIANCE MATRIX OF S-CS-6 PRINCIPAL COMPONENT SERIES

			S-6		CS-6		
		t <sub>1</sub>	ι,	٤3	<sup>n</sup> 1	72	
	4.	362299.490	9.116	18.085	162180.856	5816.690	
S-6	4,2	9.116	5540.858	073	-5460.831	1113.406	
	53	18.085	073	3085.627	10.105	11.309	
		162180.856	-5460.831	10.105	99600.481	854	
CS-6	n2	5816.690	1113.406	11.309	854	3819.884	

TABLE 35 - CORRELATION MATRIX OF N-CN-4 PRINCIPAL COMPONENT SERIES

			1	1-4		CN-4			
		41	¢ 2	٤3	54	"1	η2	<sup>n</sup> 3	<sup>n</sup> 4
	6,	1.000	000	000	000	.800	189	077	152
N-4	42	000	1.000	000	000	.185	.022	135	107
19-4	52	000	000	1.000	.000	-,309	.250	.239	092
	54	000	000	-000	1.000	.100	294	180	.027
	71	.800	.185	309	.100	1.000	000	.000	000
CN-4	n2	189	.022	.250	294	000	1.000	.000	000
CH-4	n3	077	135	. 239	180	.000	.000	1.000	.000
	n4	152	107	092	.027	000	000	.000	1.000

TABLE 36 - CORRELATION MATRIX OF N-CN-6 PRINCIPAL COMPONENT SERIES

			1	N-6			CN-6			
		41	42	€3	64	91	<sup>n</sup> 2	<sup>n</sup> 3	74	
	ξ,	1.000	.000	.001	.000	.773	218	215	091	
N-6	62	.000	1.000	.001	.000	.151	.091	111	089	
44-0	50	001	.001	1.000	.000	358	.129	.004	031	
	64	.000	.000	.000	1.000	.166	303	214	.044	
	71	.773	.151	358	.166	1.000	.000	.000	000	
CN-6	72	-,218	.091	.129	303	.000	1.000	000	000	
C14-6	03	215	111	.004	214	.000	000	1.000	000	
	74	091	089	031	.044	000	000	000	1.000	

TABLE 37 - CORRELATION MATRIX OF S-CS-4
PRINCIPAL COMPONENT SERIES

			S-4		CS-4		
		εI	£2	€3	<sup>n</sup> 1	٦2	
	€1	1.000	000	001	.851	168	
S-4	£ 2	000	1.00	000	237	208	
	€3	001	-,000	1.000	017	024	
CS-4	71	.851	237	017	1.000	.000	
	n <sub>2</sub>	168	208	024	.000	1.000	

TABLE 38 - CORRELATION MATRIX OF S-CS-6 PRINCIPAL COMPONENT SERIES

		41	¢ 2	٤3	<b>†</b> 1	<sup>n</sup> 2
	A. 171-1-1		S-6		C	S-6
	41	1.000	.000	.001	.854	.156
8-6	42	.000	1.000	000	232	. 242
	(3	.001	000	1.000	.001	.003
CS-6	72	.854	232	.001	1.000	000
	<sup>n</sup> 2	,156	.242	.003	000	1.000

5.2 The minimum number of years needed to detect a 10% increase in runoff based on the principal components. The minimum number of years, N\*, for detecting the increase of one-tenth of the old runoff means can be computed from equation (1) again,

$$N^* = \frac{\tau^2}{\mu' V^{-1} \mu}$$
 (24)

where  $\tau^2$  = the noncentrality parameter,

 $\underline{\mu} = \underline{\mu}^* - \underline{\mu}_0 ,$ 

 $\underline{\underline{\mu}}^*$  = the mean vector of the runoff variables for the seeded period,

 $\underline{\mu}_{o}$  = the mean vector of the runoff variables for the period before seeding, and

 $\underline{v}^{-1}$  = the inverse of covariance matrix  $\underline{v}$  .

The values of  $\tau^2$  are given in Table 1.

With this table the number of years needed to detect the increase can be computed easily. The values of  $N^*$  are shown in Table 39.

5.3 The application of canonical analysis. In this analysis the set of the runoff variables in the target region is first combined with the set of those in the control region. As for the principal component analysis, the computation of the canonical variables were performed on the CDC 6400 digital computer of the University Computer Center at Colorado State University. The steps in finding the coefficients for the canonical variables were described in Chapter III Section 3.

After the coefficients of the canonical variables for N-4, N-6, CN-4, CN-6, S-4, S-6, CS-4 and CS-6 are all computed and tabulated in Tables 40-47, the canonical series of each region are easily calculated [12].

TABLE 39 - MINIMUM NUMBER OF YEARS TO DETECT THE INCREASE OF 10 PERCENT IN RUNOFF MEAN USING PRINCIPAL COMPONENTS

Туре	No. of principal components in target	No. of principal components in control	Value of <u>µ'V</u> -1	τ 2	Minimum numb of years t detect the increase, N	to e
N-CN-4	4	4	1.066	11.655	11	The minimum value of N*
N-CN-6	4	4	0.813	11.655	15	is obtained from the
S-CS-4	3	2	0.243	8.640	36	larger of $N^* = \tau^2/\underline{\mu} \cdot \underline{V}^{-1}\underline{\mu}$
S-CS-6	3	2	0.273	8.640	32	or $N^* = k + 1$ where k is the total number of
	ç <sup>a</sup>				3	components in both tar- get and control regions

TABLE	40 -	COEFFICIENTS	FOR	THE	CANONICAL	VARIABLES	OF	N-4	
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CSU Sta. No.	USGS Sta. No.	lst Variable	2nd Variable	3rd Variable	4th Variable
1970000	9.0105	003956	006628	003592	.018543
1960000	9.0110	.003128	009783	.011935	042535
1866000	9.0165	.005767	004685	.026278	.009310
1830000	9.0190	.000796	.003972	002342	.002199
1820000	9.0155	001320	002450	001804	.001937
1802730	9.0265	.008752	.008348	.024461	012694
1776000	9.0360	.008385	.002618	023413	.019209

TABLE 41 - COEFFICIENTS FOR THE CANONICAL VARIABLES OF N-6

CSU Sta. No.	USGS Sta. No.	lst Variable	2nd Variable	3rd Variable	4th Variable
1970000	9.0105	006033	009805	005114	.032451
1960000	9.0110	.004802	007516	.011799	033462
1866000	9.0165	.009597	.005297	.041092	001293
1830000	9.0190	.001003	.003991	004721	.002069
1820000	9.0195	001910	003885	001016	.002457
1802730	9.0265	.013825	.025201	.014417	.035857
1776000	9.0360	.010705	008078	021553	008330

TABLE 42 - COEFFICIENTS FOR THE CANONICAL VARIABLES OF CN-4

CSU Sta. No.	USGS Sta. No.	1st Variable	2nd Variable	3rd Variable	4th Variable
1742100	9.0535	.001564	.001900	.003294	.003207
1740000	9.0575	.000620	000087	002110	002931
1720000	9.0595	.000086	001640	004363	.002621
1666300	9.0645	001139	.015690	.001530	.004480
1594260	9.0780	.001374	.001985	002575	.003694
1594236	9.0785	003596	040136	.019047	.013573
1590000	9.0850	.000525	.000354	001849	001949
1379000	9.1090	.002959	.029446	003503	.007790
1378400	9.1100	004647	030526	.005096	010398
1378100	9.1125	.001847	.006202	.009424	.002551
1377825	9.1135	.001334	.011723	.005682	002092
1377500	9.1145	000174	.000685	001047	.000878
1377280	9.1155	003380	010015	008777	003425
1377230	9.1180	.008358	.033933	.021453	.033986

TABLE 43 - COEFFICIENTS FOR THE CANONICAL VARIABLES OF CN-6

CSU Sta. No.	USGS Sta. No.	lst Variable	2nd Variable	3rd Variable	4th Variable
1742100	9.0535	.002365	001262	.006282	.000515
1740000	9.0575	.000555	000926	004814	.000489
1720000	9.0595	.000399	003421	005800	000807
1666300	9.0645	002081	.019407	009901	.012097
1594260	9.0780	.002055	.004230	001362	.003426
1594236	9.0785	003831	049581	.040428	.029201
1590000	9.0850	.000478	000866	002470	002815
1379000	9.1090	.006095	.041344	006513	013108
1378400	9.1100	008394	045125	.001749	003420
1378100	9.1125	.004031	.013690	.010848	.005814
1377825	9.1135	.001566	.017428	000181	005258
1377500	9.1145	000219	.000494	001734	.001375
1377280	9.1155	005293	011148	003944	.002038
1377230	9.1180	.013811	.053074	.011299	.031200

TABLE 44 - COEFFICIENTS FOR THE CANONICAL VARIABLES OF S-4

CSU Sta. No.	USGS Sta. No.	lst Variable	2nd Variable	3rd Variable	4th Variable
1278800	9.1650	000949	.010797	004734	.000415
1278050	9.1665	.002273	002086	.003651	002148
1272445	9.1725	.000895	.002056	.008422	004012
1077090	9.3440	.000256	.009180	.003705	009945
1073480	9.3575	.007460	.009551	023825	047496
1073436	9.3615	003435	003435	002076	.008596

TABLE 45 - COEFFICIENTS FOR THE CANONICAL VARIABLES OF S-6

CSU Sta. No.	USGS Sta. No.	lst Variable	2nd Variable	3rd Variable	4th Variable
1278800	9.1650	001790	.017228	.001080	003766
1278050	9.1665	.003301	004374	.004401	006282
1272445	9.1725	.001264	000937	.010721	014541
1077090	9.3440	.000707	.007274	011061	.011509
1073480	9.3575	.014087	.018233	052315	056559
1073436	9.3615	001675	002813	.000808	.013618

TABLE 46.- COEFFICIENTS FOR THE CANONICAL VARIABLES OF CS-4

CSU Sta. No.	USGS Sta. No.	lst Variable	2nd Variable	3rd Variable	4th Variable
1425625	9.0975	.001734	004611	.004333	005155
1377280	9.1155	003347	.055553	.035314	034470
1377230	9.1180	005054	.005968	.000726	.005480
1377200	9.1190	.003365	014545	017608	.013770
1373900	9.1275	.000054	003457	.002488	02948
1373055	9.1325	.000225	.000372	000186	000379
1373020	9.1345	.002328	.007410	007410	022485
1371815	9.1430	.004076	.010501	.000507	.023824
1370300	9.1520	.010696	.012852	.036629	.024040

TABLE 47 - COEFFICIENTS FOR THE CANONICAL VARIABLES OF CS-6

CSU Sta. No.	USGS Sta. No.	lst Variable	2nd Variable	3rd Variable	4th Variable
1425625	9.0975	.003565	010123	.004390	000819
1377280	9.1155	007498	.056512	.054781	081378
1377230	9.1180	006890	.014796	000303	.016373
1377200	9,1190	.005324	012051	020296	.025777
1373900	9.1275	.004037	053190	028107	030019
1373055	9.1325	.000329	.001040	.000236	003660
1373020	9.1345	.004450	000957	031094	.005098
1371815	9.1430	.005299	.012212	.024251	.036120
1370300	9.1520	.010325	.000287	.053491	.006772

The series of the canonical variables are tabulated in Tables 48-55 for N-4, N-6, CN-4, CN-6, S-4, S-6, CS-4, and CS-6, respectively. The means and standard deviations of the canonical series are shown in Table 56.

TABLE 48 - N-4 CANONICAL SERIES (cfs)

Year	ς <sub>1</sub>	<sup>5</sup> 2	<sup>ζ</sup> 3	ζ <sub>4</sub>
1938	5.356	-2.195	-1.045	.827
1939	2.963	-2.067	-1.608	1.737
1940	2.621	-1.618	.602	1.452
1941	4.203	-1.884	3.477	2.474
1942	4.103	582	923	5.512
1943	3.128	-2.357	364	1.256
1944	2.971	-2.235	.303	1.703
1945	3.402	-1.408	306	2.703
1946	3.167	-1.729	.509	4.003
1947	5.015	291	-1.002	2.790
1948	3.835	.538	802	1.186
1949	3.636	-3.039	.383	3.314
1950	3.739	.635	523	2.882
1951	4.328	-2.584	.609	2.509
1952	4.910	-2.824	.700	2.171
1953	2.966	-2.111	.030	3.728
1954	1.655	943	.533	2.180
1955	2.723	476	.674	2.865
1956	3.098	-1.871	248	2.906
1957	4.865	-2.033	.654	3.208
1958	3.365	-2.401	.150	3.345
1959	2.979	-1.403	.070	2.835
1960	2.855	-3.081	-1.656	3.486
1961	2.184	-2.026	791	2.903
1962	3.978	-3.188	-1.473	1.454
1963	1.504	-1.735	.921	2.158
1964	2.148	-2.119	679	3.336
1965	3,473	-3.180	720	3.962
1966	1.651	-1.613	401	1.842
1967	2.634	-2.746	196	2.709

TABLE 49 - N-6 CANONICAL SERIES (cfs)

Year	ζ1	ζ2	ζ3	<sup>ζ</sup> 4
1938	5.220	-2.656	.059	505
1939	2.893	-2.476	902	1.870
1940	2.679	-1.385	1.301	.978
1941	4.411	711	4.330	.885
1942	4.138	805	152	3.364
1943	3.198	-2.354	.342	1.723
1944	3.012	-1.995	1.019	1.780
1945	3.747	~1.514	.188	2.513
1946	3.236	-1.534	1.364	2.704
1947	5.197	650	825	1.721
1948	3.884	.209	482	298
1949	3.712	-2.711	1.024	3.491
1950	3.721	.334	572	1.811
1951	4.513	-2.259	1.273	2.446
1952	5.098	-2.365	1.539	1.697
1953	3.143	-1.963	.609	3.348
1954	1.736	761	.828	1.569
1955	2.933	278	.877	1.845
1956	3.152	-1.874	.550	1.478
1957	5.089	-1.963	1.283	1.943
1958	3.370	-2.207	.873	2.707
1959	3.080	-1.422	1.006	.546
1960	2.774	-3.571	238	1.683
1961	2.289	-2.217	.168	1.377
1962	4.004	-3.549	196	.046
1963	1.706	-1.470	1.505	1.636
1964	2.227	-2.279	.342	1.72
1965	3.805	-3.339	1.237	1.36
1966	1.851	-1.664	.974	.259
1967	2.803	-2.707	1.528	.89

TABLE 50 - CN-4 CANONICAL SERIES (cfs)

Year	ε1	€2	€3	€4
1938	5.491	.115	-1.004	420
1939	2.734	116	-1.548	1.356
1940	2.615	.354	.971	1.517
1941	4.068	.150	3.984	1.335
1942	4.074	2.026	968	4.727
1943	3.220	.655	.902	005
1944	3.106	109	.309	1.172
1945	3.502	1.417	.684	1.071
1946	3.261	.466	008	.919
1947	5.064	2.293	718	1.325
1948	4.185	2.310	353	.906
1949	3.901	062	.090	1.921
1950	3.628	2.314	.803	.889
1951	4.410	.050	.907	1.639
1952	4.792	.306	.297	1.187
1953	3.284	-1.026	.337	2.767
1954	1.706	1.475	.351	.867
1955	2.617	1.851	.501	2.010
1956	3.454	.852	.812	1.785
1957	5.042	562	.234	1.013
1958	3.567	-0.	1.252	1.950
1959	3.244	1.619	080	1.708
1960	2.788	-1.017	883	1.251
1961	2.161	.360	346	.701
1962	3.934	921	-1.099	.704
1963	1.605	.767	.207	1.260
1964	2.151	.491	.053	1.798
1965	3.539	812	.163	3.907
1966	1.896	.599	422	.600
1967	2.791	149	.551	1.160

TABLE 51 - CN-6 CANONICAL SERIES (cfs)

Year	ε <sub>1</sub>	ε2	€3	ε <sub>4</sub>
1938	5.435	743	753	-1.522
1939	2.761	672	-1.612	2.013
1940	2.720	.376	.907	1.868
1941	4.347	1.186	3.955	.553
1942	4.185	1.426	-1.182	3.186
1943	3.364	.413	.095	1.183
1944	3.191	404	.263	1.457
1945	3.889	.973	041	1.240
1946	3.395	020	033	.823
1947.	5.229	1.489	-1.339	.433
1948	4.290	1.883	778	139
1949	4.025	170	034	1.808
1950	3.739	2.145	209	.809
1951	4.659	.076	.437	2.087
1952	5.008	.190	012	1.158
1953	3.472	-1.173	.497	2.826
1954	1.825	1.054	.052	140
1955	2.919	1.474	004	1.060
1956	3.554	.719	.244	1.561
1957	5.326	878	.242	1.018
1958	3.617	.154	.726	2.551
1959	3.364	1.098	704	1.214
1960	2.811	-1.515	516	.645
1961	2.313	274	274	.450
1962	3.998	-1.686	719	.351
1963	1.888	.542	128	1.010
1964	2.322	.302	.169	.812
1965	3.921	-1.187	1.097	1.002
1966	2.035	.174	273	195
1967	3.040	127	1.315	522

TABLE 52 - S-4 CANONICAL SERIES (cfs)

TABLE 53 - S-6 CANONICAL SERIES (cfs)

Year	ζ1	<sup>ζ</sup> 2	ζ3	ζ4	
1938	4.095	.364	-2.671	-4.194	
1939	2.327	.148	980	-5.317	
1940	2.453	015	-1.387	-4.918	
1941	4.406	2.663	.354	-2.341	
1942	4.391	427	-1.419	-4.347	
1943	2.809	2.182	-2.294	-3.036	
1944	3.989	096	-1.276	-1.624	
1945	2.834	3.730	-2.042	-3.576	
1946	2.414	203	-1.678	-4.181	
1947	3.075	.861	-3.444	-4.148	
1948	3.735	.191	-2.762	-3.410	
1949	3.437	.397	-4.331	-1.361	
1950	2.313	805	805	-2.759	
1951	1.656	1.645	-2.952	-4.037	
1952	4.550	1.260	-3.162	-3.385	
1953	1.937	1.592	-1.501	-3.254	
1954	1.450	.972	-2.460	-2.073	
1955	1.786	1.169	-1.523	-2.373	
1956	2.116	1.677	-2.476	-4.210	
1957	4.638	2.586	-2.328	-5.720	
1958	3.719	1.944	092	-2.539	
1959	1.478	1.392	-2.539	-4.031	
1960	2.794	1.582	-2.113	-3.544	
1961	2.280	1.493	-1.499	-3.458	
1962	2.926	1.492	-2.984	-3.647	
1963	1.642	1.144	-2.166	-3.056	
1964	1.992	1.696	-1.201	-3.911	
1965	3.494	1.786	-3.343	-2.664	
1966	2.378	1.268	-2.392	-3.432	
1967	1.625	1.464	-1.932	-3.333	

Year	<sup>ζ</sup> 1	ς <sub>2</sub>	ζ3	ζ4
1938	4.315	.650	-3.827	-1.785
1939	2.608	295	-3.195	-3.121
1940	2.712	147	-3.104	-3.009
1941	4.493	1.887	200	-1.089
1942	4.659	473	-3.052	-1.967
1943	3.037	2.460	-2.033	-2.540
1944	4.106	.041	-1.266	-1.033
1945	2.979	3.682	-1.863	-2.466
1946	2.701	155	-2.611	-3.445
1947	3.413	1.457	-3.637	-3.310
1948	3.950	.576	-3.158	-2.331
1949	3.596	1.477	-3.813	.509
1950	2.501	872	-1.819	-1.651
1951	1.887	2.023	-3.754	-2.471
1952	4.765	1.748	-3.567	-1.583
1953	2.149	1.476	-2.031	-2.269
1954	1.611	1.328	-2.774	413
1955	1.982	1.232	-1.707	-1.486
1956	2.338	1.780	-3.107	-2.676
1957	5.055	2.684	-3.450	-4.024
1958	3.792	1.383	.019	-2.808
1959	1.737	1.697	-3.562	-2.468
1960	2.923	1.542	-2.718	-1.512
1961	2.436	1.360	-2.011	-2.307
1962	3.145	1.739	-3.473	-1.878
1963	1.881	1.324	-2.507	-1.724
1964	2.215	1.561	-1.735	-3.363
1965	3.766	2.294	-3.333	807
1966	2.578	1.285	-3.051	899
1967	1.901	1.538	-2.819	-1.276

TABLE 54 - CS-4 CANONICAL SERIES (cfs)

TABLE 55 - CS-6 CANONICAL SERIES (cfs)

V	100			
Year	ε1	ε2	ε3	€4
1938	2.930	1.215	1.510	-3.400
1939	1.591	2.091	1.492	-1.727
1940	1.759	.584	3.668	-2.275
1941	3.394	3.652	3.261	218
1942	3.803	.792	1.554	-1.400
1943	1.592	2.963	.687	-1.611
1944	3.313	1.341	2.705	-1.182
1945	2.203	4.083	2.156	240
1946	1.248	1.090	1.964	-1.471
1947	2.324	2.780	.819	605
1948	2.743	.375	.301	867
1949	2.928	1.488	815	.334
1950	1.582	1.640	.697	-1.436
1951	1.172	3.505	1.263	-1.684
1952	3.543	3.331	.329	868
1953	1.218	3.747	1.844	-1.956
1954	.790	1.778	1.233	007
1955	1.242	2.389	1.687	296
1956	1.121	2.417	.685	-1.332
1957	4.062	3.149	353	-3.382
1958	3.096	2.103	2.758	973
1959	.921	1.808	1.273	-3.346
1960	1.737	3.652	1.126	-2.610
1961	.962	2.936	1.963	-1.602
1962	2.434	2.865	1.149	-2.768
1963	.634	1.824	.478	-1.287
1964	1.480	2.019	1.262	-2.769
1965	2.462	2.433	.434	-1.807
1966	1.234	1.464	.990	835
1967	.924	1.732	.216	-1.053

Year	ε1	ε2	€3	ε <sub>4</sub>	
1938	3.125	.666	268	-1.690	
1939	1.867	1.119	.241	796	
1940	1.871	579	1.762	-2.090	
1941	3.368	2.396	2.874	.131	
1942	3.969	.358	.766	335	
1943	1.786	2.652	.400	-1.213	
1944	3.375	.546	2.335	-1.515	
1945	2.292	3.207	2.501	102	
1946	1.463	.475	1.185	-1.301	
1947	2.513	2.392	.927	.182	
1948	2.948	.382	.056	645	
1949	3.056	1.832	294	.970	
1950	1.764	1.431	.319	692	
1951	1.245	3.109	1.424	-2.270	
1952	3.640	3.039	.844	329	
1953	1.387	3.068	1.520	-2.353	
1954	.848	1.630	1.126	.792	
1955	1.357	2.024	1.420	.655	
1956	1.223	2.196	.605	-1.273	
1957	4.357	3.251	-1.240	-2.029	
1958	3.165	1.284	2.821	-1.556	
1959	1.145	1.309	355	-1.342	
1960	1.855	2.945	.362	880	
1961	1.025	2.263	1.682	-1.794	
1962	2.618	2.358	.671	-2.190	
1963	.816	1.432	026	750	
1964	1.645	1.454	.473	-2.581	
1965	2.718	2.316	082	-1.514	
1966	1.418	1.063	.549	018	
1967	1.119	1.632	258	.303	

TABLE 56 - MEANS AND STANDARD DEVIATIONS OF CANONICAL VARIABLES

Туре	Canonical Variable	Mean (cfs)	Std. Dev (cfs)
	ζ.	3.315	1.000
N-4 series	r.1	-1.819	1.000
	r2	-0.104	1,000
	51 52 53 54	2.648	1.000
		3.421	1.000
N-6 series	r.1	-1.804	1.000
	72	.695	1.000
	51 52 53 54	1.620	1.000
	ε,	3,394	1.000
CN-4 series	£2	0.523	1,000
	ε-2	0,199	1.000
	ε1 ε2 ε3 ε4	1.434	1.000
	ε,	3.555	1.000
CN-6 series	€2	,227	1.000
	€.	.046	1.000
	€1 €2 €3 €4	1.020	1.000
	5,	2.825	1.000
S-4 series	52	1.172	1.000
	5,	-2.047	1.000
	51 52 53 54	-3.463	1.000
	5,	3.041	1.000
S-6 series	52	1.276	1.000
	57	-2.639	1.000
	51 52 53 54	-2.040	1.000
	ε,	2.015	1.000
CS-4 series	ε2	2.241	1.000
	EZ	1.278	1.000
	ε1 ε2 ε3	-1.489	1.000
	ε,	2.166	1,000
CS-6 series	ε2	1.775	1.000
	€ 2	.811	1.000
	ε1 ε2 ε3 ε4	-0.941	1.000

Similar to the principal component analysis, it is clear now that,

$$E\{\zeta_i^{\star}\} = (1+h) E\{\zeta_i\}$$

where 100h is the percent increase of the runoff means in the target region. If h = 0.10, then,

$$E\{\zeta_{i}^{*}\} = 1.1 E\{\zeta_{i}\}$$

and

$$E\{\varepsilon_i^*\} = E\{\varepsilon_i\}$$

where  $\epsilon_i^*$  is the i<sup>th</sup> canonical variable of the control region for the seeded period.

The covariance matrices of N-CN-4, N-CN-6, S-CS-4, and S-CS-6 are shown in Tables 57-60, respectively. In this analysis the correlation matrices are the same as the covariance matrices since all the canonical variables have unit variances.

5.4 The minimum number of years needed to detect a 10% increase in runoff based on the canonical variables. As discussed before in Section 5.2, the minimum number of years needed to detect the increase can be obtained with the use of Table 1, which gives the value of  $\tau^2$ . After the canonical analysis has been performed because the high corre-

TABLE 57 - COVARIANCE MATRIX OF N-CN-4 CANONICAL SERIES

			N	-4	CN-4				
		¢1	¢ 2	¢3	64	۴1	* <sub>2</sub>	٤3	*4
N-4	61	1.000	0.	0.	0.	.989	0.	0.	0.
	42	0.	1.000	0.	0.	0.	.890	0.	0.
	53	0.	0.	1.000	0.	0.	0.	.847	0.
	C4	0.	0.	0.	1.000	0.	0.	0.	.767
	41	.989	0.	0.	0.	1.000	0.	0.	0.
	• 2	0.	.890	0.	0.	0.	1.000	0.	0.
CN-4	63	0.	0.	.847	0.	0.	0.	1.000	0.
	4	0.	0.	0.	.767	0.	0.	0.	1.000

TABLE 58 - COVARIANCE MATRIX OF N-CN-6 CANONICAL SERIES

			N	-6		CN-4				
		41	¢2	¢3	64	¢1	¢ 2	¢3	*4	
N-6 62	š.,	1.000	0.	0.	0.	.990	0.	0.	0.	
		0.	1.000	0.	0.	0.	.894	0.	0.	
	53	0.	0.	1.000	0.	0.	0.	.869	0.	
	54	0.	0.	0.	1.000	0.	0.	0.	.768	
	£1	.990	0.	0.	0.	1.000	0.	0.	0.	
CN-6	6.2	0.	.894	0.	0.	0.	1.000	0.	0.	
		0.	0.	. 869	0.	0.	0.	1.000	0.	
	c4	0.	0.	0.	.768	0.	0.	0.	1.000	

TABLE 59 - COVARIANCE MATRIX OF S-CS-4 CANONICAL SERIES

			S	-4		CS-4				
		۲,	¢ 2	43	4	¢1	£ 2	£ 3	° 4	
	6,	1.000	0.	0.	0.	.968	0.	0.	0.	
S-4 5	£ 2	0.	1.000	0.	0.	0.	.771	0.	0.	
	52	0.	0.	1.000	0.	0.	0.	.703	0.	
	54	0.	0.	0.	1.000	0.	0.	0.	.617	
	6,	. 968	0.	0.	0.	1.000	0.	0.	0.	
CS-4	£2	0.	.771	0.	0.	0.	1.000	0.	0.	
	43	0.	0.	.703	0.	0.	0.	1.000	0.	
	64	0.	0.	0.	.617	0.	0.	٥.	1.000	

TABLE 60 - COVARIANCE MATRIX OF S-CS-6 CANONICAL SERIES

			S	-6		CS-6				
		¢1	¢2	63	54	¢1	€2	£3	£ 4	
	51	1.000	0.	0.	0.	.969	0.	0.	0.	
	6,	0.	1.000	0.	0.	0.	.777	0.	0.	
S-6	62	0.	0.	1.000	0.	0.	0.	.696	0.	
	4	0.	0.	0.	1.000	0.	0.	0.	.568	
	E 2	.969	0.	0.	0.	1.000	0.	0.	0.	
CS-6	62	0.	.777	0.	0.	0.	1.000	0.	0.	
	63	0.	0.	.696	0.	0.	0.	1.000	0.	
	4	0.	0.	0.	.568	0.	0.	0.	1.000	

lation between target and control variables are desirable here, only the highly correlated canonical variables will be retained for further study.

For example, consider the case of S-CS-4. The correlation between the first canonical variable in S-4 and the first canonical variable in CS-4 is found to be 0.968, which is the maximum of all the correlations between the canonical variables for S-CS-4. If it is decided to use only these two canonical variables in the test, then all one needs to do is the following. From Table 56, obtain

$$\underline{\mu}_{\mathbf{o}} = \begin{bmatrix} 2.825 \\ 2.015 \end{bmatrix}$$

Assuming that there is an increase of 10% in the means of the target region and the means in the control region remain unchanged, then, the mean vector for the seeding period can be obtained as

$$\underline{\mu}^{\star} = \begin{bmatrix} 3.107 \\ 2.015 \end{bmatrix}.$$

Now  $\underline{\mu} = (\underline{\mu}^* - \underline{\mu}_0)$ , that is,

$$\underline{\nu} = \begin{bmatrix} 3.107 \\ 2.015 \end{bmatrix} - \begin{bmatrix} 2.825 \\ 2.015 \end{bmatrix}$$

$$\underline{\mu} = \begin{bmatrix} 0.282 \\ 0.0 \end{bmatrix} .$$

Compute the inverse of the covariance matrix of the first canonical variables in the target and control regions,  $\underline{V}^{-1}$ . In this case,

$$\underline{\mathbf{v}}^{-1} = \begin{bmatrix} 15.879 & -15.371 \\ \\ -15.371 & 15.879 \end{bmatrix}$$

and then compute,

$$\underline{\mu} \cdot \underline{V}^{-1} \underline{\mu} = \begin{bmatrix} 0.282 & 0.0 \end{bmatrix} \begin{bmatrix} 15.879 & -15.371 \\ -15.371 & 15.879 \end{bmatrix} \begin{bmatrix} 0.282 \\ 0.0 \end{bmatrix}$$
$$= 1.271 .$$

TABLE 61 - INVERSE OF COVARIANCE MATRIX OF N-CN-4 CANONICAL SERIES

			N	-4		CN-4				
		· ¢1	¢2	۲3	¢4	41	¢ 2	c 3	°4	
	¢ 1	45.706	0.	0.	٥.	-45.203	0.	٥.	0.	
N-4	£ 2	0.	4.810	0.	0.	0.	-4.281	0.	0.	
	63	0.	0.	3.539	0.	0.	0.	-2.997	0.	
		0.	0.	0.	2.429	0.	0.	0.	-1.863	
	.,	-45.203	0.	0.	0.	45.706	0.	0.	0.	
	.,	0.	-4.281	0.	0.	0.	4.810	0.	0.	
CN-4	63	0.	0.	-2.997	0.	0.	0.	3.539	0.	
	E 4	0.	0.	0.	-1.863	0.	0.	0.	2.429	

TABLE 62 - INVERSE OF COVARIANCE MATRIX OF N-CN-6 CANONICAL SERIES

			N-6				CN-6				
		ς,	ς2	٤3	4	۴1	¢ 2	٤3	4		
	¢,	50.251	0.	0.	0.	-49.749	0.	0.	0.		
N-6	¢ 2	0.	4.981	0.	0.	0.	-4.453	0.	0.		
100	43	0.	0.	4.084	0.	0.	0.	-3.549	0.		
	54	0.	0.	0.	2.438	0.	0.	0.	-1.872		
	£ 1	-49.749	0.	0.	0.	50.251	0.	0.	0.		
CN-6	£ 2	0.	-4.453	0.	0.	0.	4.981	0.	0.		
	63	0.	0.	-3.549	0.	0.	0.	4.084	0.		
	£ 4	0.	0.	0.	-1.872	0.	0.	0.	2.438		

TABLE 63 - INVERSE OF COVARIANCE MATRIX OF S-CS-4 CANONICAL SERIES

		s-4				CS-4				
		ζ1	ς2	٤3	4	٤1	£ 2	د,	€4	
	41	15.879	0.	0.	0.	-15.371	0.	0.	0.	
S-4	52	0.	2.466	0.	0.	0.	-1.901	0.	0.	
	63	0.	a.	1.977	0.	0.	0.	-1.390	0.	
	44	0.	0.	0.	1.615	0.	0.	0.	996	
	£ 1	-15.371	0.	0.	0.	15.879	0.	0.	0.	
CS-4	€2	0.	-1.901	0.	0.	0.	2.466	0.	0.	
	63	0.	٥.	-1.390	0.	0.	0.	1.977	0.	
	64	0.	0.	0.	996	0.	0.	0.	1.615	

TABLE 64 - INVERSE OF COVARIANCE MATRIX OF S-CS-6 CANONICAL SERIES

		s-6				CS-6				
		¢1	¢2	۲3	4	¢1	E 2	٤3	E 4	
	41	16.383	0.	0.	0.	-15.875	٥.	0.	0.	
S-6	52	0.	2.524	0.	0.	0.	-1.961	0.	0.	
	63	0.	0.	1.940	0.	0.	0.	-1.350	0.	
	4	0.	0.	0.	1.476	0.	0.	0.	839	
	•1	-15.875	0.	0.	0.	16.383	0.	0.	0.	
CS-6	£2	0.	-1.961	0.	0.	0.	2.524	0.	0.	
	63	0.	0.	-1.350	0.	0.	0.	1.940	0.	
		0.	0.	0.	839	0.	0.	0.	1.476	

The degrees of freedom here are 2 and 28, which are the number of canonical variables and the number of observations less the number of canonical variables, respectively. With these degrees of freedom, the value of  $\tau^2$  is found to be 5.468, at the level of significance  $\alpha$  = 0.05 and power  $\beta$  = 0.50. Now from

$$N^* = \frac{\tau^2}{\mu' V^{-1} \mu}$$

the value of N\* is obtained as

$$N^* = \frac{5.468}{1.271} = 4.3 = 5 \text{ years}$$
,

since  $N^*$  must be an integer. These values of  $N^*$  are shown in Table 65.

The previous results are based on the assumption that the sample mean is the same as the population mean during the non-seeded period. Now consider what effect a violation of this assumption would have on the results.

Suppose the true population mean is not equal to the sample mean. Instead it lies at the upper extremity of the 50% confidence interval established for the sample mean of the non-seeded period. Then a 10% increase in the true population mean results in a larger absolute increase than does a 10% increase in the assumed population mean (simply because the actual population mean is larger than the assumed population mean).

In the northern region, an actual 10% increase in the true population mean yields a 14.2% increase in the assumed population mean. This results in a reduction in the number of observations required to detect a change. The number of observations would be reduced to 50% of the previously determined number of observations. Similarly, in the southern region an

TABLE 65 - MINIMUM NUMBER OF YEARS TO DETECT THE INCREASE OF 10 PERCENT IN RUNOFF MEANS USING CANONICAL VARIABLES

Туре	No. of canonical variables in target	No. of canonical variables in control	Value of <u>u'V</u> -1 <u>u</u>	τ 2	Minimum number of years to detect the increase, N*	Remarks
	1	1	5.037	5.468	3	The minimum value of N*
N-CN-4	2	2	5.197	7.640	5	PER CONTENTS OF TWO CASES
	3	3	5.198	9.646	3 5 7	is obtained from the
	4	4	5.368	11.655	9	larger of N*= $\tau^2/\underline{\nu}'\underline{v}^{-1}$
	1	1	5.877	5.468	3	or $N \neq k + 1$ where k
N-CN-6	2	2	6.040	7.640	3 5	is the total number of
	3	3	6.060	9.646	7	97 NO DE 197 198 199 1
	4	1 2 3 4	6.124	11.655	ģ	variables in both targe
						and control
	1	1	1.271	5.468	5	
S-CS-4	2	1 2 3 4	1.305	7.640	5 6 7	
	3	3	1.388	9.646	7	
	4	4	1.581	11.655	9	
	1	1	1.423	5.468	4	
S-CS-6	2	1 2 3	1.465	7.640	6	
	3	3	1.690	9.646	7	
	4	4	1.752	11.655	9	

actual 10% increase in the true population mean yields a 15.6% increase in the assumed population mean, and a corresponding reduction in the required number of observations by 60 percent.

Now, suppose that the true population mean lies at the lower end of the 50% confidence interval. Then a 10% increase in the true population mean results in a smaller absolute increase than does a 10% increase in the assumed population mean.

In the northern region, an actual 10% increase in the true population mean yields a 5.8% increase in the assumed population mean. This results in an

increase in the number of observations required to detect a change. The number of observations would be increased by a factor of three. Similarly, in the southern region an actual 10% increase in the true population mean yields a 4.4% increase in the assumed population mean, and the number of observations required would be increased by a factor of 5.2.

In view of the above discussion, it is seen that if the number of observations is calculated by assuming different values for the population mean a distribution is obtained. The median number of observations will be the same as that number obtained by using the sample mean of the non-seeded period.

### Chapter VI

### CONCLUSIONS

It was the objective of this study to develop a technique for detection of a geographically widespread change in a minimum amount of time.

It was found that a combination of techniques, namely canonical analysis and multivariate  $T^2$  test was the most effective means to provide positive results in the least time. Assuming a 10% increase in runoff, 3 and 4 years are the minimum number of years needed for significance in the Upper Basin of the Colorado and the San Juan Mountains, respectively.

A word of caution is needed at this point. If the effect of precipitation management is to produce exactly a uniform 10% increase in runoff the use of only one set of canonical components is very efficient, However, if the increase is not uniform, it is safer to use several canonical components. With more canonical components, however, the number of years needed for significance increases.

It is apparent that there exists a trade-off between power of the test and representativity of the tested variables. This is well illustrated by the combined use of principal components analysis and the  ${\bf T}^2$  test. The first three or four principal

components account for 99% of the total variation in the target regions. These sets of components so to speak, are 99% representative. The number of years calculated from the  $T^2$  test is much higher than the corresponding figure for the same number of canonical components. This number of years could be decreased by using only one principal component, which already accounts on the average for 90% of the total variation. (This number was not actually calculated but the validity of the statement can be inferred from examination of the covariance matrices).

Note that when the  $\chi^2$ -test is applied to each target station with the best correlated control station, the lowest minimum number of years is found to be seven in both northern and southern regions. Again, a single station is, of course, poorly representative of the entire region. The technique (canonical components -  $T^2$  test) improves both the power of the test and the regional representativity of the tested variable, over what it would have been even with the best single target control pair.

The results from the use of four-months or six-months spring runoff are very similar. Nevertheless, better results are obtained with the six-months runoff series, particularly in the southern area.

# LIST OF SYMBOLS

Symbol	Meaning
$\mathbf{Q_i}$	Runoff at station i (i is the number in the 'Seq. No.' column in Table 2)
${\bf q_i}$	Observation of Q <sub>i</sub>
$\overline{q}_i$	The mean of Q <sub>i</sub>
q <sub>i,m</sub>	The $m^{th}$ observation of $Q_i$
2	Column vector of runoff at all stations
$\mathbf{q_i}$	Column vector of the i <sup>th</sup> observation of $Q$
<u>a</u>	Mean vector of observations of $\underline{Q}$
N	Number of observations of non-seeded period
N*	Minimum number of years for detecting a 10% increase in the runoff means of seeded period
N-4	Four-month runoff series in the northern target region (the 4 months are: April, May, June, and July)
N-6	Six-month runoff series in the northern target region (the 6 months are: March, April, May, June, July and August)
CN-4	Four-month runoff series in the northern control region
CN-6	Six-month runoff series in the northern control region
S-4	Four-month runoff series in the southern target region
S-6	Six-month runoff series in the southern target region
CS-4	Four-month runoff series in the southern control region
CS-6	Six-month runoff series in the southern control region
N-CN-4	The combination of N-4 and CN-4
N-CN-6	The combination of N-6 and CN-6
S-CS-4	The combination of S-4 and CS-4
S-CS-6	The combination of S-6 and CS-6
k	Total number of runoff variables, i.e., the number of all target and control variables
h	The fractional increase in the runoff mean
E{}	The expected value of {}
p	The number of runoff variables in target (or control) region in the principal component analysis
$p_1$	The number of runoff variables in target region
P <sub>2</sub>	The number of runoff variables in control region
$\frac{\beta}{1}$	Column vector of coefficients for computing the $i^{\mbox{th}}$ principal component

# LIST OF SYMBOLS - Continued

Symbol	Meaning
$^{\beta}$ ij	Coefficient of runoff at station j in the computation of the $\mathbf{i}^{\mbox{th}}$ principal component
I	Identity matrix
<u>v</u>	Covariance matrix of runoff variables
<u>v</u> <u>v</u> <sup>-1</sup> <u>w</u>	Inverse of V
<u>w</u>	<u>v</u> -1
ξ <sub>i</sub>	The ith principal component of target region before seeding
ξį	The i <sup>th</sup> principal component of target region for seeded period
'ni	The ith principal component of control region before seeding
n <b>*</b>	The ith principal component of control region for seeded period
ξ <sub>i,m</sub>	The mth data point of &
η <sub>i,m</sub>	The $m^{th}$ data point of $n_i$
$^{\lambda}{}_{\mathbf{i}}$	The amount of variance accounted for by the $i^{\mbox{th}}$ principal component
ζ <sub>i</sub>	The i <sup>th</sup> canonical variable of target region before seeding
ζ <b>*</b>	The $i^{\mbox{th}}$ canonical variable of target region for seeded period
$\epsilon_{ extbf{i}}$	The i <sup>th</sup> canonical variable of control region before seeding
$\epsilon_{f i}^{m{\star}}$	The i <sup>th</sup> canonical variable of control region for the seeded period
ζ <sub>i,m</sub>	The $m^{th}$ data point of $\zeta_i$
ε <sub>i,m</sub>	The $m^{\text{th}}$ data point of $\epsilon_{i}$
θ <sub>i</sub>	Correlation between $\zeta_i$ and $\epsilon_i$
<u>~</u> i	Vector of coefficients for computing $\boldsymbol{\varsigma}_{i}$
$\underline{\Upsilon}_{\mathbf{i}}$	Vector of coefficients for computing $ \epsilon_{ \underline{i}} $
$\alpha_{i,j}$	Coefficient of runoff at station j (target region) in the computation of $\boldsymbol{\varsigma}_{i}$
Y <sub>i,j</sub>	Coefficient of runoff at station j (control region) in the computation of $\boldsymbol{\epsilon}_i$
<u>μ</u> *	Runoff mean vector for the seeded period
$\frac{\mu}{\mathbf{o}}$	Runoff mean vector for the non-seeded period
<u>µ</u>	$\underline{\nu}^{\star} - \underline{\nu}_{o}$
<u>u</u> *	Transpose of $\underline{\mu}$
<u>⊬</u> ' N ∑ i=1	Summation from i=1 to i=N
Ν π i=1	Product from i=1 to i=N
τ2	Noncentrality parameter
•	Estimated value

## LIST OF SYMBOLS - Continued

-	Symbol	Meaning
	•	Transpose of a matrix
	$\sigma_{\mathtt{i}\mathtt{i}}$	Variance of runoff variable Q
	$\sigma_{\mathtt{ij}}$	Covariance of runoff variables $Q_i$ and $Q_j$
	*	Of seeded period
	cfs	Cubic feet per second

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Key Words: Statistical discrimination, regional hydrologic change, seasonal runoff, precipitation management, evaluation

Abstract: The object of this study is to find answers to the following questions: What is the appropriate statistical test for a regional target-control technique of evaluation? What is a suitable method for reduction of an originally large number of variables? Which of the Upper Basin of the Colorado River or the San Juan Mountains is a more suitable area of operations, if the effectiveness of precipitation management is to be detected as quickly as possible? The results of this research study show: 1. The T2-test is the appropriate test for multiple target-control technique of evaluation. 2. The canonical anlysis is the suitable method for the reduction of a large number of original variables. 3. The Upper Basin of the Colorado River is preferable under the assumption of an equal percentage of increase in runoff. However, if the percentage increase in the southern area is at least 1.2 times as large as in the northern area (and recent publications suggest that this ratio is probably around 3) then the southern area is preferable. Based on the T2-test, the minimum number of years for detecting an increase of 10 percent in spring runoff means are three years in the Upper Basin of the Colorado River, and four years in the San Juan Mountains.

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